Ambiguity without a State Space

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Abstract

Many decisions involve both imprecise probabilities and intractable states of the world. Objective expected utility assumes unambiguous probabilities; subjective expected utility assumes a completely specified state space. This paper analyzes a third domain of preference: sets of consequential lotteries. Using this domain, we develop a theory of Knightian ambiguity without explicitly invoking any state space. We characterize a representation that integrates a monotone transformation of first order expected utility with respect to a second order measure. The concavity of the transformation and the weighting of the measure capture ambiguity aversion. We propose a definition for comparative ambiguity aversion and uniquely characterize absolute ambiguity neutrality.

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1 Introduction

Consider a terminally ill patient whose doctor suggests two treatments. The first is an established pharmaceutical. Numerous published studies concur that this drug is successful in thirty percent of cases. The second is a new experimental surgery. Its preliminary trials suggest a success rate between twenty and forty-five percent. The two treatments are mutually exclusive, so the patient must choose between them. Can we help the patient by framing her problem with either the von Neumann–Morgenstern (1944) or Savage (1954) theory of choice under uncertainty?

We cannot frame the patient's problem in the standard von Neumann–Morgenstern (henceforth vNM) setting; the surgery is associated with an ambiguous range of possible success rates. This deficiency in the primitives, the tacit assumption of precise probability, cannot be salvaged by relaxing axioms. The subjective theory of Savage (1954) does not assume the probabilities of different outcomes are exogenously precise.¹ On the other hand, to invoke the Savage machinery, our patient must be able to: first, determine the relevant states of nature; second, decide how each choice assigns consequences to these states.

She fails both counts and cannot frame her problem using subjective utility theory. Regarding the state space, the patient has no medical training and does not understand what the relevant states are. She knows only the information presented by her physician, expressed entirely in the space of probabilities over consequences. Even given a comprehensive list of states, the designers of the experimental surgery are unsure which states would make the surgery more likely to be successful. Studies are needed exactly because the mapping from states to outcomes that actually represents this new medical procedure is still unknown. More generally, an inability to correctly formulate the state space or the acts is often the cause of the ambiguity in a decision problem. As with objective vNM theory, the deficiency is not any particular subjective utility representation, but is fundamental to the structure of an act.

The machinery of vNM is too simple to express the patient's problem; the machinery of Savage is too complicated. This paper studies an alternative framework to analyze such decisions under ambiguity: sets of lotteries over consequences. This domain incorporates ambiguity without appealing to any state space, thus avoiding the technology of states and acts. Each set captures the possible distributions on consequences associated with a particular option. By enriching the domain of preference, we can introduce ambiguity in an objective setting. We can then express the patient's decision problem in formal terms: the established drug is represented as the singleton lottery that yields success with probability 0.3 and the experimental surgery as the set of lotteries that yield success with probabilities between 0.2 and 0.45.

Even without understanding the states underlying her choices, the decision maker may still understand how her choices affect consequences, which are the ultimate objects of her utility. She can understand how an option might make her feel, without understanding the causal mechanism or act that delivers that feeling. Without access to a state space, the agent forms some boundaries

¹An important variation is the model of Anscombe and Aumann (1963), where the consequences are a mixture space, which we subsume in our discussion of general subjective Savage theory.

on the possible consequential probabilities associated with each choice. These restrictions are captured as sets of lotteries. Such restrictions on the space of consequential uncertainty seem especially plausible when summary information is given to the decision maker by an expert, like the doctor in our introduction.

The standard subjective approach does not assume exogenous ambiguity. Insofar as ambiguity exists, it is meaningful only in the mind of the agent. This austere view makes no additional assumptions of the world outside the agent's mind. While such parsimony is theoretically elegant, we believe there are compelling reasons to allow objective ambiguity.

Introducing objective ambiguity as a set of lotteries arms us with more detail, and this detail can capture realistic features of the decision problem. Fully subjective theories provide no device for the agent to incorporate outside information, or lack thereof, about uncertainty into her decision making. For example, the information provided by her doctor comprises an important part of our decision maker's problem.

Subjective utility theory also has difficulty distinguishing situations without any ambiguity and situations where the decision maker resolves ambiguity by selecting a single probability. For example, the maxmin expected utility model (Gilboa and Schmeidler 1989) represents attitude towards ambiguity by a set of multiple priors. The same set of priors also represents the existence of ambiguity. This confounds the presence of ambiguity with the agent's resolution of that ambiguity. If the agent is unsophisticated, we cannot distinguish whether ambiguity or some other factor causing her irrational assessment. If she obeys the Savage axioms, we cannot distinguish whether ambiguity was resolved in a probabilistically sophisticated manner or there was never any ambiguity to be resolved in the first place. These features are difficult to separate in a purely subjective model, where the existence of and attitude towards ambiguity are identified simultaneously, but are easily separated in a model where the the existence of ambiguity is exogenous and independent of its resolution. By starting with assumed sets of lotteries, our model forces this separation by fiat.

Roughly speaking, we propose the following utility representation U(A) for a set A of lotteries:

$$U(A) = \frac{\int_A \phi \circ u \, d\mu}{\mu(A)},$$

where u is a standard affine expected utility function on single lotteries, μ is a probability measure on sets of lotteries, and ϕ is an increasing transformation applied to u. This decision maker considers all of the relevant lotteries in A when making her decision, and their relative consideration is fixed across sets by a measure over all lotteries. Her attitude to ambiguity is captured by the transformation ϕ and her weighting μ .

This paper contributes to a recent literature which equips the decision problem with exogenous ambiguity through sets of lotteries or priors. Jaffray (1989) first introduced exogenous ambiguity over lotteries by defining preferences over non-additive belief functions. He imposes the mixture space axioms and characterizes a generalization of expected utility. More recently, two papers, independent of and contemporaneous with this one, also use sets of objective lotteries to model ambiguity, but arrive at distinct utility representations. Olszewski (2003) extends the axioms of Dekel (1986), most notably betweenness and dominance, to sets of lotteries and characterizes a generalized form of α -maxmin utility, which evaluates a set by a convex combination of its minimal and maximal elements and is further discussed in Section 2. Stinchcombe (2003) presents a novel dual formulation of the Expected Utility Theorem on the mixture space of sets, invoking the standard Archimedean and independence conditions. An advantage of these linear approaches is that the resulting representations are well defined over lower dimensional sets, which our measure theoretic approach must finesse. Finally, Ghirardato (2001), Jaffray and Jeleva (2004), and Nehring (1999) consider multi-valued acts which map states of the world to sets of consequences.

Another recent strand of research enriches the subjective model with exogenous information about the priors on the state space. Gajdos, Tallon, and Vergnaud (2004a) and Hayashi (2003) consider preferences defined on act-set pairs, where the set captures the possible probabilities on the state space. Gajdos, Tallon, and Vergnaud (2004b) and Wang (2003) further enrich the problem with reference or anchor priors, and consider preference over act-set-anchor triples. A possible technical reconciliation between our approach and these is a generalized form of probabilistic sophistication, where an ambiguous act is evaluated by its induced set of distributions over consequences. This reconciliation and its limitations are further discussed when the model is formally introduced in Section 2. These models take the set of lotteries or priors as exogenous; Ghirardato, Maccheroni, and Marinacci (2004), Nehring (2001), and Siniscalchi (forthcoming) suggest various behavioral methods to identify the perceived priors in subjective models.

Defining preference over sets has a rich tradition in the economic literature of choice under ignorance, starting with Arrow and Hurwicz (1972). There, the decision maker chooses between finite sets of sure consequences, but has no further control as to which alternatives are eventually selected from these sets. This interpretation is closely related to ours, since the decision maker knows only that some set of objects or lotteries is possible and has no further information on how nature will select a particular object in the set. On the other hand, Kreps (1979) uses sets of sure consequences as a way of identifying a preference for flexibility. He interprets as menus; the decision maker will eventually choose an available option from the menu. A desire for larger menus suggests a desire to keep one's options open.

More recently, Dekel, Lipman, and Rustichini (2001) and Gul and Pesendorfer (2001) use menus of lotteries to model flexibility and commitment. However, their interpretation of sets is fundamentally different from ours. They view sets as menus of stochastic choices; at an implicit second stage, the agent chooses a single lottery from the menu. They extend the interpretation over sets due to Kreps (1979); we extend the interpretation due to Arrow and Hurwicz (1972). For us, a set reflects objective information about the risks involved in a decision—there is no implicit second stage of choice. The set of lotteries represents the possible risks associated with the first stage of choice. For example, in the Ellsberg urn, the subject knows that a bet on yellow is associated with a range of lotteries, but has no further information beyond that range. She certainly cannot select the distribution of colors in the urn at some later stage. The utility representations between these papers and ours are not nested; it is straightforward to construct utility functions that satisfy our axioms but violate axioms in either Dekel, Lipman, and Rustichini (2001) or Gul and Pesendorfer (2001), and vice versa. Most notably, the independence conditions of these two papers are not imposed in our main representation.

Dekel, Lipman, and Rustichini do share our reservations with Savage's state space, in the context of unforeseen contingencies. Using menus of lotteries, they identify subjective states which capture perceived taste uncertainty over final outcomes. Our interpretation of sets as informative boundaries on risk is more blunt, but we feel it speaks more directly to ambiguity, while Dekel, Lipman, and Rustichini speak more directly to flexibility and unforeseen contingencies. Their decision maker is additively separable or probabilistically sophisticated over her subjectively comprehensive list of possible tastes, precluding ambiguity.² In the end, we feel that the appropriate interpretation of sets, either as menus or as information, depends on the specific application at hand; a broad view of models and interpretations seems in keeping with the motivations of both papers.

Our representation has formal and interpretive antecedents. A version of a general representation for conditional expectation was already proven and applied to Jeffrey's (1965) syntactic theory of decision by Bolker (1966, 1967), who deserves original credit for the mathematical result.³ While the mathematical domain of Bolker's result is somewhat different than ours, the substantive assumptions are essentially similar. The technical differences between Bolker's result and ours will be explained more comprehensively when we present our theorem.

Our interpretation of the mathematical result differs from Jeffrey's theory of joint desirability and probability over logical propositions or sentences, which is not a treatment of ambiguity. While seemingly established in formal philosophy, Jeffrey's framework is less familiar in economics. It does not immediately translate to standard formulations, as it is unclear whether Jeffrey's propositions should be interpreted as consequences, as states, or as acts.⁴ Jeffrey's probability on propositions is the only mechanic carrier of uncertainty in the model and is not allowed to vary, whereas we consider the space of *all* possible lotteries on consequences. Perhaps our theory can be viewed as a form of Jeffrey's that considers "propositions" regarding risk: both theories have preferences for information, captured as a sentence about the world or as a set of lotteries, and both implement some form of conditional expectation. That said, we suspect Jeffrey might object to our basic model and primitives: "I take it to be the principal virtue of the present theory, that it makes no use of the notion of a gamble or of any other causal notion (Jeffrey 1965, p. 147)."

Our expression of ambiguity aversion through nonlinear aggregation of expected utilities is closely related to the literature which links ambiguity aversion to second order risk aversion. Segal (1987, 1990) first modeled ambiguity using objective two-stage lotteries and pioneered the conceptualization of ambiguity aversion as a failure to reduce compound lotteries. More recently, Segal's

 $^{^{2}}$ A recent paper by Epstein and Marinacci (2005) allows a set of priors on the subjective states which is resolved through maxmin expected utility.

 $^{^{3}}$ We are extremely indebted to Larry Epstein for bringing Bolker's work to our attention, and thank Chris Chambers for subsequent references on the Jeffrey model.

⁴Nonetheless, there is at least one application of the theory to welfare economics (Broome 1990).

approach has inspired extensions to subjective settings (Ergin and Gul 2004, Klibanoff, Marinacci, and Mukerji 2005, Nau 2003), where first and second order uncertainties are modeled as the two dimensions of a product state space. These papers' representations also feature transformations of expected utility functions. The spirit of Segal's approach is also present in Halevy and Feltkamp (2005), who propose that ambiguous prospects may be perceived as bundles of correlated risky prospects. Halevy (2004) presents experimental evidence that suggests a connection between ambiguity aversion and reduction of compound lotteries; he also provides a nice categorization and comparative testing of the different proposed forms of compounding.

While our representation and its associated interpretation are complementary, our domain is distinct from these studies. We examine *sets* of lotteries, not two-stage *lotteries* over lotteries. Referring back to the Ellsberg urn, for example, the subject's given information is perhaps more transparently modeled as a set of possible distributions of colors ("between zero and sixty yellow balls") than as a compound process of distributions over distributions of colors.

We include no explicit or verifiable second order uncertainty of any sort; the decision maker is not asked to rank bets on which lottery actually generates risk. Since they are outside our purview, we have nothing directly to say about reduction of compound lotteries. This is arguably a comparative virtue: the uncertainty over lotteries is produced without appeal to second order measurement devices. For example, we do not allow bets on the measure over lotteries or over which lottery actually obtains, and cannot elicit information about ambiguity by isolating and varying outcomes on which lottery represents the fundamental uncertainty. Insofar as "second order" uncertainty exists, it is as an artifact of the suggested utility representation, rather than as a primitive assumption of the model's domain. This benefit comes at a cost. The use of direct second order measurement identifies the nonlinear aggregator ϕ modulo positive affine transformations with two parameters of freedom. Dropping this direct measurement and working with sets of lotteries, we cannot achieve this level of uniqueness and are forced to allow a third degree of freedom.

In the next section, we formally introduce the primitives of our theory. Section 3 contains our main representation: the decision maker integrates a transformed expected utility with respect to a second order measure, conditioning on the objective set of lotteries. Section 4 discusses comparative and absolute ambiguity aversion and conducts some comparative statics.

2 An objective model of ambiguity

We introduce the domain of preference, a special family of sets of objective lotteries. Preferences over somewhat different familiees of sets of lotteries are also studied by Olszewski (2003), who considers convex polyhedra, and Stinchcombe (2003), who consider closed convex sets. Neither family is a superset or subset of the family we introduce here.

The finite set X denotes the set of deterministic outcomes. ΔX is the set of lotteries on X, endowed with the topology of weak convergence, which is induced by the Euclidean metric when ΔX is represented as $\Delta X = \{x \in \mathbb{R}^{|X|-1}_+ : \sum_{i=1}^{|X|-1} x_i \leq 1\}$. $\Delta^2 X = \Delta(\Delta X)$ is the set of Borel probability measures on ΔX , or second order measures on X. We refer to elements of ΔX as "lotteries," and reserve the term "measures" for elements of $\Delta^2 X$.

A set is regular is it is equal to the closure of its interior. Our domain of choice is the family of nonempty regular and singleton subsets of ΔX , denoted

$$\mathcal{K}^*(\Delta X) = \{A \subseteq \Delta X : \overline{\operatorname{int}(A)} = A \text{ or } |A| = 1\} \setminus \emptyset.$$

This domain restriction is a significant one. It excludes, for example, all finite subsets with more than one element. The decision maker must either face ambiguity regarding the probabilities of all consequences or face no ambiguity at all; she cannot know precisely the probabilities of some outcomes but not of others.⁵ While regular sets are dense in the family of closed subsets, a weakness of our utility representation is that conditional expectation is not defined on null sets. We could include such sets in our domain, but our measure-theoretic representation would have no bite on them.⁶ Their ability to put structure on such sets is a comparative strength of the linear approaches of Olszewski (2003) and Stinchcombe (2003).

Any singleton is unambiguous, because the risk is known and precise. We view this simple definition as a strength of the theory, but also acknowledge it is an artifact of the exogenous nature of ambiguity in the model.

We maintain throughout that \succeq is a complete and transitive binary relation on $\mathcal{K}^*(\Delta X)$, with \succ and \sim having the standard definitions.

Given a probability assessment μ over states, each act f is naturally associated with its induced distribution over consequences: $\nu = \mu \circ f^{-1}$. If the decision maker is probabilistically sophisticated, in the sense of Machina and Schmeidler (1992), these image lotteries completely characterize her preference. Then an act contains more structure than is required for decision making; all payoff relevant information is captured by its distribution. Similarly, the sets of consequential lotteries might be viewed as reduced formulations of ambiguity in a Savage setting if the relevant information is captured by the *set* of possible distributions induced by an act, given a set of probability assessments over states. For example, in the Ellsberg urn, such a reduction implies that the agent treats yellow and black symmetrically and is indifferent between betting on either. This imposed symmetry seems reasonable in many cases. Ellsberg himself reported, "In our examples, actual subjects do tend to be indifferent between betting on [yellow or black]. . . . the reasons, if any, to favor one or the other balanced out subjectively so that the possibilities entered into their final decisions weighted equivalently (Ellsberg 1961, p. 658)."

In other cases, the translation is more tenuous. The following example due to Hayashi (2003). Consider two urns: the first contains 100 red or green balls, the second contains 100 red, green, or yellow balls. No further information regarding these urns is known. If the decision maker uses

⁵Another way of thinking about the restriction, pointed out to us by Wojciech Olszewski, is that these sets are of the same dimension as ΔX .

⁶Another facile alternative would be to impose the following axioms separately on lower dimensional components of ΔX , for example each face or edge of ΔX . While this would provide structure on comparing two subsets of the same face, it would not restrict preference across faces.

only the induced distributions of acts, then betting on a red ball from the first urn induces the same lotteries as betting on a red ball from the second urn, yet she may prefer to bet on the first urn because there are fewer possible colors. This suggests that some important information might be lost in reducing ambiguous acts to their distributions. The association of ambiguous acts in subjective settings to ambiguous sets in our objective model is a delicate one. A relaxation of distributional reduction is proposed by Gajdos, Tallon, and Vergnaud (2004a).

Finally, this model sharply delineates ambiguity as a closed set of lotteries. In reality, the decision maker may not have such crisp boundaries on the possible lotteries. Instead, she may think a variety of sets may represent the actual ambiguity, and have a belief on the likelihood of these sets, an element of $\Delta(\mathcal{K}(\Delta X))$. Moreover, if the agent can hold ambiguous beliefs about the consequences ΔX , then she may also hold ambiguous beliefs about the ambiguity, captured as $\mathcal{K}(\Delta(\mathcal{K}(\Delta X)))$. Iterative applications of risk $\Delta(\cdot)$ and ambiguity collapse to a single expanded space of consequences, then the model loses no generality if the consequence space is properly constructed. We prove the hypothesis in another paper, which constructs a universal type space of ambiguous beliefs that also provides a universal consequence space for this model (Ahn 2003).

3 Representation

One possible resolution of ambiguity is to focus on the worst possible lottery. Let $u : \Delta X \to \mathbb{R}$ be a utility function on single lotteries and consider the following utility function for a set A:

$$U(A) = \min_{a \in A} u(a).$$

This translates the seminal representation of Gilboa and Schmeidler (1989) to our setting.

Maxmin utility has a clean functional form and crisp axiomatic characterizations. Nonetheless, aside from the minimal lottery, the objective form of this representation ignores all the other lotteries included in a set A. Indeeed, Ellsberg anticipated with dissatisfaction: "In almost no cases . . . will the *only* fact worth noting about a prospective action be its 'security level': the 'worst' of the expectations associated with reasonably possible probability distributions. To choose on a 'maxmin' criterion alone would be to ignore entirely those probability judgments for which there is evidence (Ellsberg 1961, p. 662)."

Partly to mitigate this extreme form of ambiguity aversion, α -maxmin utility takes a weighted combination of the worst and best distributions in a set, and is characterized by Ghirardato, Maccheroni, and Marinacci (2004) in a subjective setting and by Olszewski (2003) in an objective setting similar to ours. While α -maxmin utility improves simple maxmin, it retains some problems. For example, α -maxmin utility still ignores almost all of the information contained in the set of priors or the set of lotteries; preferences are completely characterized by minimal and maximal elements. Take the lotteries over \$0 and \$100, represented on [0,1] by their probabilities for \$100. Then α -maxmin utility is indifferent between $[0, 0.5] \cup [0.9, 1]$ and $[0, 0.1] \cup [0.5, 1]$, while the latter might be more intuitively appealing. For an example involving convex sets, consider a decision maker facing three consequences $X = \{a, b, c\}$ and evaluates sets using α -maxmin expected utility with index u(a) = 8, u(b) = 4, and u(c) = 0 and with $\alpha = 3/4$. Her utility for the set $A = \{x \in \Delta X : x(b) \leq 1/2\}$, the lotteries where b is less likely that not, is $\frac{3}{4} \cdot 0 + \frac{1}{4} \cdot 8 = 2$. Her utility for the set $B = \{x \in \Delta X : x(b) \geq 1/2\}$, the lotteries where b is more likely than not, is $\frac{3}{4} \cdot 2 + \frac{1}{4} \cdot 6 = 3$. Hence B is strictly preferred to A. Her utility for their union ΔX is 2, so she is indifferent between A and ΔX , seemingly ignoring the good information, relative to A, embedded in ΔX with B.

We suggest an alternative representation. Consider a utility function u on the single lotteries and a probability measure μ on the Borel subsets of ΔX . We propose the following resolution of ambiguity:

$$U(A) = \frac{\int_A u \, d\mu}{\mu(A)}.$$

The agent conditions her utility on a second order measure, given the information that the set A of lotteries obtains. This incorporates every lottery in A, weighted by the measure μ .

Ideally, the utility function should be continuous and the measure μ should be nonatomic. Hausdorff continuity of the preference would provide both conditions. Unfortunately, the proposed utility violates continuity when defined over all closed sets of lotteries. Suppose the lotteries are over two outcomes, winning \$0 or \$100. This set of lotteries can be represented as [0, 1], indexed by the probability of winning \$100. Set u(x) = x and μ to the Lebesgue measure. Then the sets $A_{\delta} = [\frac{1}{4} - \delta, \frac{1}{4} + \delta] \cup [\frac{3}{4} - 2\delta, \frac{3}{4} + 2\delta]$ and $B_{\delta} = [\frac{1}{4} - 2\delta, \frac{1}{4} + 2\delta] \cup [\frac{3}{4} - \delta, \frac{3}{4} + \delta]$ both converge to the doubleton $\{\frac{1}{4}, \frac{3}{4}\}$ as $\delta \to 0$. Yet $U(A_{\delta}) = \frac{7}{12}$ and $U(B_{\delta}) = \frac{5}{12}$ for all δ , so continuity fails. The main problem is that $\{\frac{1}{4}, \frac{3}{4}\}$ is a Lebesgue null set, so its conditional expectation is undefined.

This motivates our construction of \mathcal{K}^* , which limits attention to regular sets and singletons and excludes pathological sets like $\{\frac{1}{4}, \frac{3}{4}\}$. Of course, we include the singletons to retain unambiguous choices. The decision maker must have either ambiguous marginals across all dimensions or else face no ambiguity at all; she cannot know precisely the probabilities of some outcomes but not of others.⁷ Throughout this section, we will restrict \gtrsim to be a binary relation on \mathcal{K}^* , rather than \mathcal{K} .⁸

Our representation involves measures, so we need closure under standard set theoretic operations. Regular sets are not closed under these operations: consider the intersection of two regular sets that meet only at their boundaries. Regular sets are closed under the regularized set operations,

⁷Another way of thinking about the restriction, pointed out to us by Wojciech Olszewski, is that regular sets are of the same dimension as ΔX .

⁸Rather than refining the domain of choice, a different approach might invoke a lexicographic probability system (Blume, Brandenburger, and Dekel 1991) of measures. Such an approach would be technically involved, since the number of null sets is very large.

defined as:

$$A \cup' B = \overline{\operatorname{int}(A \cup B)};$$

$$A \cap' B = \overline{\operatorname{int}(A \cap B)};$$

$$A \setminus' B = \overline{\operatorname{int}(A \setminus B)}.$$

We slightly abuse notation and drop the primes: all subsequent set operations are regularized. We can now introduce the axioms.

Axiom 1 (Downward Hausdorff continuity). Suppose A_n converges to $\{a\}$ in the Hausdorff metric. $A_n \succeq B$ for all n implies $\{a\} \succeq B$ and $A_n \preceq B$ for all n implies $\{a\} \preceq B$.

This axiom is somewhat weaker than the standard Hausdorff continuity condition, which assumes the implication for all Hausdorff convergent sequences of sets. Here, it is assumed only for sequences which are decreasing to a point.

We now introduce some notation. Let λ denote the Lebesgue measure on ΔX . If μ is absolutely continuous with respect to Lebesgue measure, this will be denoted $\mu \ll \lambda$. Symmetric set difference is noted by $A \triangle B = (A \cup B) \setminus (A \cap B) = (A \setminus B) \cup (B \setminus A)$. The next condition asserts that, as the measure of the difference of two sets converges to zero, the sets also converge in preference.

Axiom 2 (Lebesgue continuity). Suppose $\lambda(A) > 0$ and $\lambda(A_n \triangle A) \to 0$. $A_n \succeq B$ for all n implies $A \succeq B$ and $B \succeq A_n$ for all n implies $B \succeq A$.⁹

Lebesgue continuity is different than Hausdorff continuity; there are sequences of sets which converge in Lebesgue measure but do not converge in Hausdorff distance, and vice versa. Since our representation is measure-theoretic in nature, this topology is more appropriate for our purposes. Olszewski (2003) does not require any form of continuity, and argues that Hausdorff continuity is questionable in the context of ambiguity, and offers an example, similar to the one given when we motivated our domain restriction, of a sequence of finite sets which might plausibly violate continuity. Our domain, \mathcal{K}^* , excludes such pathologies. This is not offered as a justification the original restriction to regular sets, but only to argue that Lebesgue continuity is reasonable given that restriction.

Axiom 3 (Disjoint set betweenness). Suppose A, B are regular and disjoint. $A \succeq B$ implies $A \succeq A \cup B \succeq B$ and $A \succ B$ implies $A \succ A \cup B \succ B$.

Gul and Pesendorfer (2001) assume a similar axiom in their work on temptation and self control. Our axiom is technically weaker in one sense, applying only to disjoint unions; it is stronger in another, preserving both weak and strict preference. More importantly, set betweenness carries a distinct substantive interpretation in our model. Gul and Pesendorfer think of set betweenness in the context of temptation and menus: unchosen or suboptimal elements of a menu may carry a

⁹Lebesgue continuity replaces a divisibility axiom which was assumed in an earlier version of this paper.

disutility of temptation. Our sets are not menus, but provide information about possible lotteries. In the context of temptation, set betweenness relaxes the following modularity condition: $A \succeq B$ implies $A \sim A \cup B$, which is satisfied by a decision maker who exhibits no preference for commitment. In the context of ambiguity, set betweenness relaxes exactly the opposite direction, implied by maxmin utility: $A \succeq B$ implies $A \cup B \sim B$.

Maxmin utility barely fails our version of set betweenness. Instead, it implies the following: if $A \succ B$, then $A \succ A \cup B \sim B$. In evaluating $A \cup B$, maxmin utility pessmistically ignores the better set A, and pays attention only to the worse set B. While maxmin utility passes the weak part of the axiom, it fails the strict part. This strict component of the axiom distinguishes our representation from maxmin utility by forcing the agent to pay attention to good news.

Similarly, α -maxmin utility also fails the strict part of the axiom. The example given earlier demonstrates this: $X = \{a, b, c\}$, $\alpha = 3/4$, u(a) = 8, u(b) = 4, and u(c) = 0. Consider $A = \{x \in \Delta X : x(b) \leq 1/2\}$ and $B = \{x \in \Delta X : x(b) \geq 1/2\}$. These sets are disjoint, in the regularized sense, and $A \succ B$, yet $A \sim A \cup B$, violating the strict part of our axiom. In words, this decision maker is ignoring the good information, relative to A, that is included in B when evaluating their union. On the other hand, like maxmin utility, α -maxmin utility does satisfy the weak form of the disjoint set betweenness. In fact, Olszewski (2003) shows that a more general class of utilities satisfies a strengthened form of weak set betweenness, which he coins "strong generalized betweenness." However, as the example demonstrates, strong generalized betweenness implies only the weak, but not the strict, component of this axiom.¹⁰

These comparisons suggest that the strict part of the axiom is particularly biting. If a strictly preferred set B of lotteries is added to the set A, then the decision maker must feel strictly better off. A failure suggests that the better component of a union of sets is irrelevant to her preference.

Axiom 4 (Balancedness). Suppose A, B, C, D are regular with $[A \cup B] \cap [C \cup D] = \emptyset$ and $A \sim B \succ C, D$ (or $A \sim B \prec C, D$). $A \cup C \succeq B \cup C$ implies $A \cup D \succeq B \cup D$.

Balancedness has the flavor of Savage's second postulate, the Sure-Thing Principle (Savage 1954, Section 2.7). Consider the sets $A \cup C$ and $B \cup C$ in the hypothesis. These two unions share C and differ only in that one has A and the other has B. Then, if we replace C with another set D which shares C's preference relation with A, this preference is preserved. The only data that matter in evaluating these types of unions are their set differences, namely A and B. Their intersection, D in the conclusion, does not affect their relative desirability. Similar intuitions, that two objects should be compared by where they are different, are provided to justify independence and monotonicity in other models. Our axiom is weaker because we add the additional restriction that $A \sim B$. Although loosely related, balancedness is *not* the same as the Sure-Thing Principle. The indifference and disjointedness assumptions really have no analog in the Savage setting. Conversely, the Sure-Thing Principle has no direct translation in our setting.

This axiom is technically similar to the classic structural property of qualitative probability: if $A \cup C$ is more likely than $B \cup C$, then $A \cup D$ is more likely than $B \cup D$, for any C, D disjoint

¹⁰The strict part of disjoint set betweenness is not always violated by maxmin utility either. For example, if

from A, B. In proving the eventual utility representation, it analogously guarantees the consistency of a likelihood relation which is directly constructed from preference. So, although on face more similar to the sure-thing principle, the technical value of balancedness is actually closer to Savage's fourth postulate, Weak Comparative Probability. The Sure-Thing Principle is generally regarded as a restriction on utility, while Weak Comparative Probability is regarded as a restriction on likelihood. Here, balancedness has connections to both, because the same space carries both utility and probability in our setting.

We note that maxmin utility is balanced in a somewhat vacuous manner: $A \cup B \sim B \cup C$ whenever $A \sim B \succ C$.

These four axioms are necessary and sufficient for the proposed utility representation.

Theorem 1. A preference relation \succeq on \mathcal{K}^* satisfies Axioms 1–4 if and only if there exist a continuous $u: \Delta X \to \mathbb{R}$ and a probability measure $\mu \ll \lambda$ on ΔX with full support such that

$$U(A) = \begin{cases} \frac{\int_A u \, d\mu}{\mu(A)} & \text{if } A \text{ is regular} \\ u(x) & \text{if } A = \{x\} \end{cases}$$

is a utility representation of \succeq .

Moreover, suppose there exists such a utility representation by (u, μ) . Then (v, ν) also represent \gtrsim if and only if

$$v(x) = \frac{au(x) + b}{cu(x) + d};$$

$$\nu(B) = \mu(B)[c\int_B u \, d\mu + d]$$

for some numbers $a, b, c, d \in \mathbb{R}$ such that ad - bc > 0 and $d = 1 - c \int_{\Delta X} u \, d\mu$.

Proof. See Appendix A.1.

Theorem 1 does not conclude that the utility function u on lotteries is affine. No form of independence is imposed, accommodating behavior under risk like the Allais paradox. To our knowledge, our allowance for nonlinear preferences over singletons is unique in the literature on preferences over sets of lotteries or priors. We consider this an important advantage of this measure-theoretic approach. Of course, adding singleton independence, as we will do shortly, provides additional structure.

As mentioned in the introduction, the technical content of the theorem rediscovers a mathematical result that was proven earlier by and should be credited to Bolker (1966), who provided a functional characterization for quotients of measures on complete nonatomic Boolean algebras. While the motivations, formal hypotheses, and proofs are different, we do not want to claim any significant technical novelty. At the same time, there are some important formal differences in the results. Bolker begins with a nonatomic algebra: this excludes the singletons and would force the existence of ambiguity in our setting. This exclusion might be natural in a propositional Boolean model for logic or syntax, but the inclusion of atoms is essential in our interpretation of sets as ambiguity to formalize unambiguous choices and to get a handle on the the form of the utility function u. To allow for atoms, we require the additional continuity assumptions. We also provide more structure on the functional equation: the Radon–Nikodym derivative, which becomes u in our setting, is continuous and μ is absolutely continuous with full support. Within this narrower class, our continuity conditions become necessary, as well as sufficient, for the representation.

While a detailed proof is in the appendix, we briefly outline the main ideas here. Our proof is presented in the case that it might be more transparent than Bolker's original arguments to decision theorists. First, let Λ_A denote the family of regular subsets of A which are indifferent to A. Λ_A is λ system, being closed under complementation and disjoint unions, both by disjoint set betweenness. Such systems are recently emphasized in the literature on ambiguity (Epstein 1999, Epstein and Zhang 2001, Zhang 1999). The balancedness axiom provides a natural likelihood relation \succeq_{ℓ} on Λ_A : $S \succeq_{\ell} T$ if and only if $S \cup B \succeq T \cup B$ for some disjoint $B \prec A$. The likelihood relation satisfies sufficient conditions, due to Zhang (1999), for the existence of a quantitative probability measure P_A which represents \succeq_l . The measure can be uniquely extended to the entire class of regular subsets indifferent to A. We then use the extended P_A to construct a signed measure ν_A on all of the regular subsets of ΔX such that the sign of $\nu_A(S)$ identifies whether an arbitrary regular set S is preferred to A: $\nu_A(S) \ge 0$ if and only if $S \succeq A$.

Each ν_A provides information about the underlying preference with respect to a fixed set A, serving as a partial representation of the preference. The next step is to connect these signed measures together to construct a complete representation. Take any three regular sets such that $A \succ B \succ C$ and consider the vector-valued measure $\nu_{ABC} = (\nu_A, \nu_B, \nu_C)$ taking values in \mathbb{R}^3 . Exploiting the constructed representation properties of the measures' signs, an application of Lyapunov Convexity Theorem demonstrates that the image of ν_{ABC} is spanned by two vectors.¹¹ This implies that ν_C is linearly determined by ν_A and ν_B . Then the family of signed measures $\mathcal{M} = \{\nu_A : A \text{ is regular}\}$ is spanned by two of its elements, ν and μ . Furthermore, the cone generated by \mathcal{M} is convex. This convexity implies that \mathcal{M} is contained in a half space of the vector space of signed measures. Let ν^* be the measure orthogonal to the boundary of this half space. Some simple algebra demonstrates that the fraction $\nu^*(A)/\mu(A)$ is a utility representation over the regular sets. Then the Radon–Nikodym Theorem implies there exists a measurable real-valued function u on ΔX such that $\nu^*(A) = \int_A u \, d\mu$. This demonstrates the utility representation for regular sets; the representation for singletons is the consequence of a technical convergence lemma and downward Hausdorff continuity.

Finally, the nonparametric domain of choice ΔX considers arbitrary distributions on X and makes no restrictive assumptions on the shape of risk. However, the proof only assumes that ΔX is a compact Polish mixture space. The characterization remains valid when restricted to a family of parameterized distributions, provided the parameter space is a compact Polish mixture space. This is true even if the consequence space X is not compact or Polish. For example, if the decision

¹¹A form of the Lyapunov Convexity Theorem was also invoked to similar effect by Bolker (1966).

maker knows that the risk is normal over all levels of wealth, we can define preference over closed subsets of a compact rectangle $M \times V \subset \mathbb{R}^2$, associated with normal distributions of different means M and variances V, while $X = \mathbb{R}$ is not compact.

As mentioned, adding independence over unambiguous singletons further refines the representation. The following is standard.

Axiom 5 (Singleton independence). For all $a, b, c \in \Delta X$ and $\alpha \in (0, 1)$, $\{a\} \succeq \{b\}$ if and only if $\{\alpha a + (1 - \alpha)c\} \succeq \{\alpha b + (1 - \alpha)c\}$.

This assumption provides a two step evaluation of a set of lotteries: first each lottery is linearly aggregated by expected utility, then the entire set is nonlinearly aggregated by μ . This is reminiscent of the compounding approach to ambiguity aversion forwarded by Uzi Segal (1987, 1990).

Corollary 2. A preference relation \succeq on \mathcal{K}^* satisfies Axioms 1–5 if and only if there exist an affine $u: \Delta X \to \mathbb{R}$, a strictly increasing and continuous $\phi: \mathbb{R} \to \mathbb{R}$, and a probability measure $\mu \ll \lambda$ on ΔX with full support such that

$$U(A) = \begin{cases} \frac{\int_A \phi \circ u \, d\mu}{\mu(A)} & \text{if } A \text{ is regular,} \\ \phi(u(x)) & \text{if } A = \{x\} \end{cases}$$

is a utility representation of \succeq .

Moreover, suppose there exists such a utility representation by (u, ϕ, μ) . Then (v, ψ, ν) also represent \succeq if and only if

$$\begin{array}{lll} v(x) &=& \alpha u(x) + \beta; \\ \psi(z) &=& \displaystyle \frac{a\phi(z) + b}{c\phi(z) + d}; \\ \nu(B) &=& \displaystyle \mu(B)[c\int_B u\,d\mu + d] \end{array}$$

for some numbers $\alpha, \beta, a, b, c, d \in \mathbb{R}$ such that $\alpha, ad - bc > 0$ and $d = 1 - c \int_{\Delta X} u \, d\mu$.

Imposing singleton independence, we retain standard expected utility as a special case on the unambiguous singletons. The utility u is a standard affine expected utility function, and ϕ is a transformation that retains the ordinal independence condition on preferences; both u and $\phi \circ u$ produce linear indifference curves on single lotteries. The classic vNM Expected Utility Theorem states there exists *some* affine utility representation, but not that *all* utility representations must be affine. There are nonlinear utility representations of independent preference, as shown in Figures 1 and 2. The classic theory is cardinal with respect to the value function on deterministic outcomes X (up to scale transformations), but is only ordinal with respect to the utility function on lotteries ΔX . A nonlinear monotone transformation is irrelevant in comparing one lottery to another, but becomes very relevant in comparing sets of lotteries. The functional form of utility used to represent preferences over single lotteries is important when we integrate that form over sets of lotteries.



Figure 1: Linear expected utility (u)



Figure 2: Nonlinear "expected" utility $(\phi \circ u)$

In Figure 1, the simplex of probability distributions ΔX lies on the ground, and the linear expected utility curve $u(\Delta X)$ floats above it. Since u represents independent preference, it has linear indifference curves. Figure 2 shows the same independent preference, except now the linear utility representation u is transformed by some concave function $\phi : \mathbb{R} \to \mathbb{R}$. Notice that the transformed utility retains the same linear indifference curves.

Although related, the curvature of utility in our model is not induced in the traditional sense of direct second order risk aversion on compound or two-stage lotteries. This would suggest a relaxation of the reduction axiom in the space of compound lotteries, about which we have nothing to say. The domain of our model is *sets* of simple lotteries, rather than *lotteries* of simple lotteries. The measure μ over lotteries is fixed by the representation; it is not assumed as a primitive of the theory, nor is it allowed to vary to reflect different second order uncertainties. Instead, the cardinality of $\phi \circ u$ is induced by the agent's preferences over subsets of ΔX .

Thus one part of an agent's attitude towards ambiguity is the transformation ϕ , which captures a cardinal intensity of preference for one type of lottery compared to another. This intensity is important because the utility integrates this intensity with respect to the fixed measure μ . The manifestation of ambiguity aversion as a departure from linearity should be comforting, since it resonates the traditional analysis of risk aversion. We develop this metaphor more carefully later.

The second part of her attitude is determined by her weighting μ . One interpretation of our setting is a game against nature, where nature decides which lottery is actually realized.¹² Then μ is the agent's belief about nature's mixed strategy. If the agent thinks μ puts more probability on worse lotteries, she thinks nature will more likely choose a bad lottery. The measure μ can be interpreted as an agent's assessment of her "luck" in ambiguous situations. The more weight she places on better lotteries, the luckier she believes herself to be. Two agents can share a common utility function u and transformation ϕ , but still have different preferences because one thinks of herself as luckier than the other. The maxmin utility is an extreme case where the agent pessimistically believes that nature always chooses the minimal element of a set.¹³ Here, our decision maker weights all the possible lotteries in a set by μ , using all of the available information.

Another interpretation of μ is less wedded to the statistical view of the decision problem as a game against nature. We can interpret μ as a measure of salience or how much attention the agent pays to the various lotteries. Worse lotteries loom larger in the minds of those with a distaste for ambiguity. The measure μ then corresponds to the personal attention given to the possible lotteries, normalized so $\mu(\Delta X) = 1$. Then μ captures the psychological, rather than statistical, weight the decision maker attaches to the various lotteries.

Similar concepts of luck cannot be formalized in a subjective model where the agent takes a fixed weighted average over her priors over states. This is because whether the weighting is optimistic or pessimistic depends on the act being evaluated. For example, if she holds a bet on yellow in the Ellsberg urn, placing more weight on distributions with many yellow balls is optimistic. On the other hand, if she holds a bet on black, the same weighting would be considered pessimistic. Only by putting the second order beliefs directly on consequences can the decision maker be considered optimistic or pessimistic.

The uniqueness of u modulo affine transformations is standard. The uniqueness of the transformation ϕ and the probability measure μ are strictly weaker than that achieved in some other

¹²This interpretation is advanced more explicitly by Olszewski (2003).

¹³Of course, this extreme sense of bad luck cannot actually be defined as a probability measure over nature's behavior.

representations. This is due to the different kind of information being elicited. Since we view the direct measurement of second order beliefs as artificial, we do not allow bets on the space ΔX and have no method to directly elicit the decision maker's second order belief. In contrast, the models of Ergin and Gul (2004) and Klibanoff, Marinacci, and Mukerji (2005) identify an exact second order measure μ on first order priors. But this identification is at least partially an artifact of these models' primitives, which ask the decision maker to assess bets on the correct prior. In contrast, we do not access this full range of bets to identify beliefs; the decision maker's probabilistic payoff under lottery x is never supplemented or changed. For example, our decision maker is not asked to compare betting a dollar on the event that the actual lottery generating outcomes is in the set A to betting on the event that the lottery is in another set B. Demanding such hypothetical comparisons would provide another measurement device that we suspect would obtain uniqueness, but one of our motivations for developing this domain is the artificiality of such contingent comparisons. We feel one of the model's strengths is the lack of such forced comparisons. We believe that deriving the second order uncertainty purely as an artifact of the utility representation, rather than embedding it into the primitives of the model, is a comparative strength. On the other hand, as mentioned in the introduction, this comes at some cost in terms of identification. Instead of linear transformations, which have two degrees of freedom, we allow fractional linear transformations with three degrees of freedom, since d is defined as a function of c. Moreover, the second order belief is not identified uniquely, but allowed a degree of freedom, for similar reasons. At a technical level, this suggests that the form of uniqueness in the second order prior model depends on the form of the elicitation.

4 Ambiguity aversion

4.1 Comparative ambiguity aversion

From here on, we take the representation of Corollary 2 as given. In this section, we develop tools to discuss ambiguity aversion in our objective setting. We first introduce concepts to compare ambiguity aversion across individuals.

Definition 1. The relation \succeq_1 is *locally more ambiguity averse at* A than \succeq_2 if

$$A \succeq_1 \{a\} \Rightarrow A \succeq_2 \{a\}$$

and

$$A \succ_1 \{a\} \Rightarrow A \succ_2 \{a\},$$

for all $a \in \Delta X$.

The relation \succeq_1 is (globally) more ambiguity averse than \succeq_2 if it is locally more ambiguity averse at all $A \in \mathcal{K}^*$.

This definition is analogous to Epstein's (1999) definition of comparative ambiguity aversion.

He considers \succeq_1 more ambiguity averse than \succeq_2 if for every arbitrary act f and any unambiguous act $g, f \succeq_1 g$ implies $f \succeq_2 g$ and $f \succ_1 g$ implies $f \succ_2 g$, where an act is considered unambiguous if it is measurable with respect to a λ -system of unambiguous events. The definition by Ghirardato and Marinacci (2002) is identical, except they further restrict g in the hypothesis to be a constant function. The two definitions disagree on what exactly constitutes an unambiguous act. Here, singleton lotteries directly replace place of λ -measurable or constant acts, allowing us to finesse the issue. With the subjective definitions proposed by Epstein and by Ghirardato and Marinacci, our definition shares the virtue of being applicable across different utility representations for the decision maker.

While the local definition of ambiguity aversion does not reference the cardinal utilities of the decision makers over X, a consequence of the global definition is that if one decision maker is globally more ambiguity averse than another, then both share the same restricted preferences on singleton lotteries. If X represents levels of wealth, if we can compare two agents' global ambiguity attitudes, the agents must have the same cardinal utility for wealth. So the ordering of ambiguity aversion is coarser than the ordering of risk aversion.

Proposition 3. If \succeq_1 is globally more ambiguity averse than \succeq_2 , then $\succeq_1|_{\Delta X} = \succeq_2|_{\Delta X}$

Proof. This follows directly from restricting the definition to singletons.

If two decision makers disagree on the desirability of uncertain prospects, it is because they have different cardinal tastes over the sure consequences or because they have different reactions to the size of ambiguity. Our definition of comparative ambiguity aversion separates the effects of risk and ambiguity on decision making by fiat. Ghirardato and Marinacci consider such separation desirable and carefully delineate conditions where it is implied by their definition in a Savage framework. The immediacy of the separation here is a mechanical consequence of our objective domain of lotteries, where we have a rich linear structure not immediately available in a Savage domain.

Given that two agents share risk preferences, we can consider a_1 and a_2 their respective ambiguity-free equivalents to A if $\{a_1\} \sim_1 A$ and $\{a_2\} \sim_2 A$.¹⁴ Ambiguity-free equivalents are conceptually similar to certainty equivalents in the theory of risk aversion. An agent is more risk averse than another if the other's monetary certainty equivalent for a lottery is greater than her own certainty equivalent. Here, we replace the natural ordering on money with the preference ordering on singletons. Then the first is more ambiguity averse at A than the second if she prefers the second's ambiguity-free equivalent a_2 to her own equivalent a_1 : $\{a_2\} \succeq_1 \{a_1\}$. She can be more ambiguity averse for some sets but less ambiguity averse for others, in the same way she can be more risk or less risk averse for different lotteries.

We mentioned in Section 3 that maxmin utility nearly meets the axioms of Theorem 1. Specifically, it only barely fails to meet disjoint set-betweenness. We now show that maxmin utility is a limit case of our representation. To consider limits in the space of possible preferences, we define

¹⁴More generally, we can consider the entire set $\{a \in \Delta X : \{a\} \sim A\}$ as the ambiguity-free equivalent. This set is the restriction of a hyperplane to ΔX if \succeq meets singleton independence.

a topology on the space of preferences. For a fixed set $A \in \mathcal{K}^*$, let

$$d_A(\succeq,\succeq') = d(\{a : \{a\} \sim A\}, \{a : \{a\} \sim' A\})$$

recalling d is the Hausdorff distance.¹⁵ We now say $\succeq_n \rightarrow \succeq$ if $d_A(\succeq_n, \succeq) \rightarrow 0$ for all $A \in \mathcal{K}^*$. So, a sequence of preferences converges if the respective ambiguity-free equivalents for A converge for any set $A \in \mathbb{Z}$.¹⁶

Remember our discussion of the two sides of ambiguity aversion: the transformation ϕ reflecting a cardinal utility towards gambles and the probability assessment μ reflecting an attitude about one's luck. To conduct comparative statics, we isolate each effect by keeping the other fixed. We begin by fixing the measure μ and comparing the curvature of ϕ . In the theory of risk, more curvature corresponds to more risk aversion. Similarly, in our theory more curvature corresponds to more ambiguity aversion. For a fixed risk profile, we let \gtrsim_{MMEU} refer to the corresponding maxmin expected utility. Given similar results in the theory of risk, the following is hardly surprising; moreover, analogous results for definitions of comparative ambiguity aversion in subjective settings are provided by Klibanoff, Marinacci, and Mukerji (2005)

Proposition 4. Suppose \succeq and \succeq' have representations (u, ϕ, μ) and (u, ϕ', μ) . Then \succeq is more ambiguity averse than \succeq' if and only if $\phi = h \circ \phi'$ for some concave and strictly increasing $h : \mathbb{R} \to \mathbb{R}$.

Moreover, suppose $\{\succeq_n\}$ have representations $\{(u, \phi_n, \mu)\}$. If each ϕ_n is twice differentiable and, for all $x \in \Delta X$,

$$\min_{x} - \frac{\phi_n''(x)}{\phi_n'(x)} \to \infty,$$

as $n \to \infty$, then $\succeq_n \to \succeq_{MMEU}$.

Proof. The first part of the proposition follows almost directly from Jensen's Inequality. For the second part, make u positive by adding a sufficiently large constant. The result follows by taking a subsequence $\phi_{n(m)}$ with $\phi_{n(m)}$ a concave transformation of $-e^{-mx}$, then taking $m \to \infty$.

The fraction $-\frac{\phi_i''(x)}{\phi_i'(x)}$ strongly resembles the standard Arrow–Pratt coefficient of absolute risk aversion. Fixing a probability μ , we can construct a similar quantitative measure to compare ambiguity aversion. As with risk, the ratio of the second to the first derivative provides a measure of curvature and captures the level of ambiguity aversion for the agent. As this measure approaches infinity, the agent's preferences get closer to maxmin utility. Her relative distaste for worse lotteries increases, and she wishes more and more to avoid sets that include such lotteries.

Now we fix the transformation ϕ and vary the probability μ . Recall μ captures the agent's perception of her luck. The partial order of stochastic dominance formalizes what it means for one agent to consider herself "luckier" than another. A measure stochastically dominates another precisely if it puts more weight on more desirable lotteries. We let $\Delta A = \{\mu \in \Delta X : \mu(A) = 1\}$ for

¹⁵The arguments in d are compact by continuity of the preferences and d_A is a semimetric on the space of preferences.

¹⁶Notice that this convergence need not be uniform across A.

any set $A \subseteq X$. Since *u* is continuous, the set of maximizers over a compact set *A* is closed. When $\mu(A) > 0$, $\mu|A$ is the conditional probability defined by $[\mu|A](S) = \frac{\mu(A \cap S)}{\mu(A)}$.¹⁷

Proposition 5. Suppose \succeq and \succeq' have representation $\{u, \phi, \mu\}$ and $\{u, \phi, \mu'\}$. If $\mu|A$ stochastically dominates $\mu'|A$ with respect to the lattice $\succeq |_A$, then \succeq is locally more ambiguity averse at A than \succeq' .¹⁸

Moreover, suppose $\{\succeq_n\}$ have representations $\{(u, \phi, \mu_n)\}$. If, for all measurable $A \subseteq \Delta X$, $\mu_n | A$ converges weakly to $\Delta(\arg\min_{x \in A} u(x))$, as $n \to \infty$, then $\succeq_n \to \succeq_{MMEU}$.

Proof. This follows directly from definitions.

This is another way maxmin utility is a limit case of our representation. The more unlucky an agent considers herself, the closer her behavior is to maxmin preference. She focuses more and more of her attention on the worse lotteries, until the worst lottery becomes the sole criterion for comparison.

4.2 Absolute ambiguity neutrality

Having established these relative definitions, we introduce an absolute benchmark for ambiguity neutrality. In the subjective literature, there is a natural benchmark of probabilistic sophistication. In our setting, probabilistic sophistication reduces the domain to the standard vNM setting of singleton lotteries. So, we should note that it is hardly obvious if "ambiguity neutrality" has any significance here. Nonetheless, there are arguably reasonable benchmarks for ambiguity neutrality. In the theory of risk aversion, risk neutrality is identified by a linearity in the Bernoulli utility function for money. Here, we also propose various types of linearity, adjusted to our special decision setting with subsets.

The first kind of linearity we impose is on the transformation ϕ . Since it measures cardinal attitudes on the space of lotteries, it can be interpreted as a measure of second order risk aversion. In other words, if ϕ really is nonlinear, then the agent might treat compound lotteries differently than their reductions. So consider the following.

Axiom 6 (Linearity). If (ϕ, u, μ) is a representation of \succeq , then $\phi(z) = \frac{az+b}{cz+d}$ for some ad - bc > 0.

This axiom essentially assumes there exists some representation where ϕ is linear. By itself, linearity of ϕ does not identify a single preference, because μ could be one of many possible measures. While this assumption is somewhat unsatisfying because it is imposed on an artifact ϕ of preference rather than on \succeq directly, this is the best that can be achieved given the lack of a second order measurement device.

We need more axioms to identify an ambiguity neutral probability assessment. While it is obvious what the ambiguity neutral transformation ϕ should be, we still need to characterize an

 $^{^{17}\}text{Since}~\mu$ has full support, this fraction is always well defined for regular sets.

¹⁸A measure μ stochastically dominates ν with respect to the lattice \gtrsim if $\int f d\mu \geq \int f d\nu$ for any function f that is monotone with respect to \succeq .

"ambiguity neutral" weighting μ . Define $\alpha A + (1 - \alpha)B = \{\alpha a + (1 - \alpha)b : a \in A, b \in B\}$. The definition of singleton independence can be expanded to include an ambiguous set on one side.

Axiom 7 (Singleton-set independence). For all $a, x \in \Delta X$, $A \subseteq \Delta X$, and $\alpha \in (0, 1)$, $\{a\} \succeq A$ if and only if $\{\alpha a + (1 - \alpha)x\} \succeq \alpha A + (1 - \alpha)\{x\}$.

The axiom states that the agent is neutral to mixtures between a set and a singleton. This is similar to the standard "hedging" axiom in subjective theories of ambiguity. The spirit of these axioms is that the agent prefers a mixture of two acts to either of the acts separately; the mixture "hedges" the ambiguity of the acts. For example, in Gilboa and Schmeidler (1989), the key axiom characterizing ambiguity aversion is that $\alpha f + (1 - \alpha)g \succeq f$ for any two indifferent acts $f \sim g$, while $\alpha f + (1 - \alpha)g \sim f$ characterizes ambiguity neutrality. Our condition is that if the agent is indifferent between an unambiguous singleton a and any set A, then she will retain that indifference when these choices are mixed with another unambiguous choice x. One feature which distinguishes our main representation is that no form of affine independence over sets is assumed at all, since none of the main axioms take any convex combinations of non-singleton sets. Such combinations are required only to characterize a special subclass of the representation.

A stronger notion of independence between sets is that $A \succeq B$ if and only if $\alpha A + (1 - \alpha)C \succeq \alpha B + (1 - \alpha)C$ for any sets A, B, C; this is invoked by Stinchcombe (2003) for sets of lotteries and by Gajdos, Tallon, and Vergnaud (2004a) for sets of priors over a state space. The justification for this stronger assumption is usually tied to the temporal resolution of uncertainty. In this interpretation, the mixture $\alpha A + (1 - \alpha)C$ is viewed as the lottery where the set A is realized with probability α and C is realized with probability $1 - \alpha$. If the decision maker is indifferent to when uncertainty resolves, she should retain her preference of A to B when they are mixed with another set C. Our representation precludes this view, since the relative likelihoods of A and C are fixed by the probability measure μ . The probabilistic interpretation of set mixtures is at odds the agent's own assessment of the relative probabilities of these sets, determined by $\mu(A)$ and $\mu(C)$. Indeed, it is easy to find examples of preferences that meet our axioms and violate set independence.

Since \succeq is assumed to be a weak order, hence complete, singleton-set independence can be expressed as the conjunction of two implications, strict and indifferent: first, if $\{a\} \succ A$, then $\{\alpha a + (1 - \alpha)x\} \succ \alpha A + (1 - \alpha)\{x\}$; second, if $\{a\} \sim A$, then $\{\alpha a + (1 - \alpha)x\} \sim \alpha A + (1 - \alpha)\{x\}$. A weaker variant of such an independence condition which only assumes the strict implication is used by Hayashi (2003) for sets of priors and noted by Gul and Pesendorfer (2001) and Dekel, Lipman, and Rustichini (2001) for menus of lotteries, with similar intuitions: for all convex $A, B, C \subseteq \Delta X$, $A \succ B$ implies $\alpha A + (1 - \alpha)C \succ \alpha B + (1 - \alpha)C$. On one hand, the required implication is weaker because it is only asserted for strict preference, and not for indifference. On the other and, we believe, more important hand, the domain of our singleton-set independence assumption is considerably smaller; the implications are imposed only when the sets A and C are singletons.

The corrected form of Weak Independence required by Dekel, Lipman, and Rustichini (2001) to derive ordinal expected utility is stronger than singleton-set independence: for all convex $A, B, C \subseteq \Delta X$: if $A \subseteq B$, then for all $\alpha \in (0, 1), A \succeq B$ if and only if $\alpha A + (1 - \alpha)C \succeq \alpha B + (1 - \alpha)C$

(Dekel, Lipman, Rustichini, and Sarver 2005). On the domain of convex subsets, singleton-set independence can be replaced with the assertion that for all $a, c \in \Delta X$ and convex $B \subseteq \Delta X$, if $a \in B$, then for all $\alpha \in (0, 1)$, $\{a\} \succeq B$ if and only if $\{\alpha a + (1 - \alpha)c\} \succeq \alpha B + (1 - \alpha)\{c\}$. This is because, given our utility representation, if A is convex, there exists some $a \in A$ such that $\{a\} \sim A$. So singleton-set independence relaxes new Weak Independence by narrowing the domain of application to A and C which are singletons.

Dekel (1986) introduced a betweenness axiom for single lotteries which weakens independence: $A \succeq B$ if and only if $A \succeq \alpha A + (1 - \alpha)B \succeq B$. This betweenness property is extended by Olszewski (2003) for sets of lotteries: if $A \succeq B$, then for any set of scalars $P \subseteq [0, 1]$, $A \succeq PA + (1 - P)B \succeq B$, where PA + (1 - P) is the union of the convex combinations of A and B by any scalar in the set P. This strong generalized betweenness property yields, as a special case, the weak form of disjoint set betweenness by setting $P = \{0, 1\}$.

A well known consequence of the Haar Measure Theorem is that Lebesgue measure is the essentially unique translation invariant measure on $\mathbb{R}^{|X|-1}$, in the sense that $\lambda(A) = \lambda(A + y)$. Lebesgue measure is a particular measure-theoretic restriction on what the decision maker would consider the "average" lottery from a set without further information. Translation invariance is arguably a "neutral" feature, since the relative weight of different sets does not vary if we make the sets uniformly better or worse. In any case, translation invariance of the measure is mathematically implied by linearity and singleton-set independence. A geometric notion of the average as the Steiner point is characterized by Gajdos, Tallon, and Vergnaud (2004a) and Hayashi (2003) for sets of priors over states and suggested by Stinchcombe (2003) for sets of lotteries over consequences.

Theorem 6. \succeq meets Axioms 6 and 7 if and only if there exists a representation (ϕ, u, μ) where ϕ is the identity function and $\mu = \lambda$.¹⁹

Proof. See Appendix A.2.

Since ambiguity neutrality is uniquely identified, we can now naturally identify ambiguity aversion with any preference which is more ambiguity averse than the representation of Theorem 6.

A Appendix

A.1 Proof of Theorem 1

Recall that λ denotes the (|X|-1)-dimensional Lebesgue probability measure on ΔX . We begin by proving a useful technical lemma. The lemma is used in the proof of Theorem 6 in Appendix A.2 and makes Corollary 2 an immediate consequence of Theorem 1.

¹⁹This is the only result in the paper that invokes the finite cardinality of X. When ΔX is infinite-dimensional, its open subsets admit no nontrivial translations. We are very grateful to Bob Anderson, Oleh Nykyforchyn, Chris Shannon, Max Stinchcombe, and especially Bill Zame for pointing out an important error in an earlier draft and many helpful conversations on its resolution.

Lemma 7. There is a version $f = \frac{d\mu}{d\lambda}$ (λ -almost everywhere) of the Radon–Nikodym derivative of μ with respect to λ that is continuous at a if and only if $\frac{\mu(A_n)}{\lambda(A_n)} \to f(a)$ whenever $A_n \to \{a\}$ in the Hausdorff metric topology and $\lambda(A_n) > 0$.

Proof. We first prove sufficiency by contradiction. Suppose $\frac{\mu(A_n)}{\lambda(A_n)} \to f(a)$ if $A_n \to \{a\}$ and f(x) is discontinuous at a for any version f. Then either max $\{f(x), f(a)\}$ or min $\{f(x), f(a)\}$ is discontinuous at a; without loss of generality assume the former. Then there exists some $\varepsilon > 0$ such that

$$D_n = \left\{ x : \|x - a\| < \frac{1}{n} \text{ and } f(a) - f(x) > \varepsilon \right\}$$

is nonempty for all n. Furthermore, $\lambda(D_n) > 0$ for all n, otherwise we can find a version g = f almost everywhere with respect to λ with g continuous at a. Then

$$\frac{\int_{D_n} f \, d\lambda}{\lambda(D_n)} < f(a) - \varepsilon$$

for all D_n while $D_n \to \{a\}$. This contradicts our assumption that $\frac{\mu(A_n)}{\lambda(A_n)} \to f(a)$ whenever $A_n \to \{a\}$.

For necessity, suppose there exists a version f that is continuous at a and take any sequence $A_n \to \{a\}$. Fix $\varepsilon > 0$. There is a corresponding $\delta > 0$ such that $|f(x) - f(a)| < \varepsilon$ whenever $||x - a|| < \delta$. As $A_n \to \{a\}$, there exists some N such that if n > N, then $||x - a|| < \delta$ for any $x \in A_n$. Then for all n > N,

$$\begin{aligned} \left| \frac{\mu(A_n)}{\lambda(A_n)} - f(a) \right| &= \left| \frac{\int_{A_n} f \, d\lambda}{\lambda(A_n)} - f(a) \right| \\ &\leq \frac{\int_{A_n} |f(x) - f(a)| \, d\lambda}{\lambda(A_n)} \\ &< \frac{\int_{A_n} \varepsilon \, d\lambda}{\lambda(A_n)} \\ &= \varepsilon. \quad \Box \end{aligned}$$

The routine verification of necessity is omitted. The uniqueness claim follows readily from Bolker (1966). We now provide a proof of sufficiency. We begin by demonstrating the existence of a utility function.

Lemma 8. If \succeq is downward Hausdorff continuous and satisfies disjoint set betweenness, then for each A, there exists a singleton $\{x\} \sim A$.

Proof. Fix an arbitrary regular set A. We first prove there exists some x^* such that $\{x^*\} \succeq A$. Let Π_n^i denote the regular partition or grid of A induced by the lattice of points whose dimensions are multiples of 2^{-n} . By repeated applications of disjoint set betweenness, at least one element B_1 of the partition Π_1 satisfies $B_1 \succeq A$. If an element B_n of Π_n satisfies $B_n \succeq A$, then a subset $B_{n+1} \subset B_n$ with $B_{n+1} \in \Pi_{n+1}$ must satisfy $B_{n+1} \succeq B_n$. Thus there exists a sequence of sets B_n such that $B_n \succeq A$ for all n. Since the elements of this sequence are decreasing and have arbitrarily small radius, they converge in the Hausdorff metric to a point $\{x^*\}$. By decreasing Hausdorff continuity, $\{x^*\} \succeq A$. Similarly, there also exists some $\{x_*\} \preceq A$. Decreasing Hausdorff continuity also implies that the restriction of \succeq to the singletons is continuous. Then there must exist some point x with $\{x^*\} \succeq \{x\} \succeq \{x_*\}$ such that $\{x\} \sim A$.

Therefore, each A has an ambiguity-free equivalent $\{x\} \sim A$. Moreover, by Debreu's Theorem, there exists a utility representation u on the singletons, which can be extended to \mathcal{K}^* through Lemma 8 by setting

 $V(A) = u(\{x_A\})$, the utility of its ambiguity-free equivalent. We extend \succeq to a larger family of sets. We will consider the family of Borel sets modulo λ .²⁰ Details of the following constructions can be found in (Halmos 1974, pp. 166–169). We write symmetric set difference as $A \triangle B = (A \setminus B) \cup (B \setminus A)$. Let \mathcal{Z}' refer to the quotient $\mathcal{B}(\Delta X)/\mathcal{N}$, where \mathcal{N} is the family of Borel sets with Lebesgue measure zero. So, if two sets differ only on a set of Lebesgue measure zero, $\lambda(A \triangle B) = 0$, then they are considered part of the same equivalence class in \mathcal{Z}' . We will remove the equivalence class of the empty set, to make $\mathcal{Z} = \mathcal{Z}' \setminus [\emptyset]$. At times, we abuse notation and refer to the equivalence class [A] by a representative element A; this should not cause any confusion. Let $\pi(A, B) = \lambda(A \triangle B)$, providing a separable metric on \mathcal{Z} Halmos (1974, Theorem B, p. 168).

Lemma 9. Let A be a Borel subset of ΔX . If $\varepsilon > 0$, there exists a regular set K such that $\lambda(K \triangle A) < \varepsilon$.

Proof. Fix a Borel set A and $\varepsilon > 0$. Since λ is an outer regular measure, there exists an open set $O \supseteq A$ such that $\lambda(O \setminus A) < \varepsilon/2$. Let \mathcal{A} denote the family of all closed cubes contained in O. This family is a Vitali covering of A.²¹ By the Vitali Covering Theorem (Dunford and Schwartz 1957, Theorem 3, p. 212), there exists a sequence of disjoint sets A_1, A_2, \ldots such that $\lambda(A \setminus \bigcup_{i=1}^{\infty} A_i) = 0$. Then there exists a finite index n such that $\lambda(A \setminus \bigcup_{i=1}^{n} A_i) < \varepsilon/2$. Let $K = \bigcup_{i=1}^{n} A_i$. K is a finite union of regular sets, hence regular. Since $K \subseteq O$, $\lambda(K \setminus A) \leq \lambda(O \setminus A) < \varepsilon/2$. Thus $\lambda(A \triangle K) = \lambda(K \setminus A) + \lambda(A \setminus K) < \varepsilon$.

Lemma 10. There exists a π -continuous extension of \succeq to \mathcal{Z} .

Proof. Fix a nonnull Borel set B. By Lemma 9, there exists a sequence A_n converging to B in Lebesgue measure measure. Let $V(B) = \lim V(A_n)$. This sequence is convergent as u is continuous and the Lebesgue measure metric is complete. The Lebesgue continuity axiom makes the selection of any particular sequence inessential. Lebesgue continuity also implies π -continuity of this extension.

Adding the singletons to \mathcal{Z} , when necessary, is done in the natural fashion. At times, we will move between the spaces with and without the singletons appended, but this should not cause any confusion.

We will prove that the axioms on the extended preference are sufficient for the utility representation on all of \mathcal{Z} . Then the utility will also represent the original \succeq on the restricted domain \mathcal{K}^* .

To prove sufficiency, we reference the theory of probability representation on λ -systems of subsets. This theory was developed conceptually in the mathematical foundations of quantum mechanics by Birkhoff and von Neumann (1936) and von Neumann (1955). More recently, such structure is exploited by Zhang (1999) and Epstein and Zhang (2001) to define the algebraic properties of unambiguous events in the Savage state space.

Definition 2. A family Λ of subsets of X is a λ -system if:

- $1. \ X \in \Lambda,$
- 2. $S \in \Lambda$ implies $S^{\complement} \in \Lambda$, and
- 3. if $A_1, A_2, \ldots \in \Lambda$ are pairwise disjoint, then $\bigcup_{n=1}^{\infty} A_n \in \Lambda$.

The family of indifferent subsets of A, notated as Λ_A , is conveniently a λ -system.

Lemma 11. The family $\Lambda_A = \{S \subseteq A : S \sim A\}$ is a λ -system (relative to A).

²⁰The particular use of Lebesgue measure here is inessential; any nonatomic measure with full support will suffice.

²¹In $\mathbb{R}^{|X|}$, a family \mathcal{A} of closed sets is a *Vitali covering* of A if each set has strictly positive Lebesgue measure and every point in A is contained in sets of \mathcal{A} with arbitrarily small diameter.

Proof. The first condition follows from completeness of \succeq . If $S \in \Lambda_A$, then its (relative) complement $A \setminus S \sim A$ to satisfy disjoint set betweenness. Closure under finite disjoint unions follows directly from disjoint set betweenness. For any countable disjoint sequence $\{A_n\}_{n=1}^{\infty}$, $\lambda(\bigcup_{n=1}^{N} A_n)$ converges to $\lambda(\bigcup_{n=1}^{\infty} A_n)$ as N goes to infinity because $\lambda(\bigcup_{n=1}^{\infty} A_n)$ is finite. Then the third condition holds by applying disjoint set betweenness to the finite unions $\bigcup_{n=1}^{N} A_n$, then passing the limit to the preferences using Lebesgue continuity.

There are various results that find sufficient conditions on a qualitative likelihood ranking \succeq_{ℓ} on a λ -system for the existence of a consistent probability measure, for example (Suppes 1966, Theorem 3) or (Krantz, Luce, Suppes, and Tversky 1971, p. 215). We state a recent version by Zhang (1999), which we find to be the most intuitive and transparent. The symbols \succ_{ℓ} and \sim_{ℓ} carry their natural meanings.

Theorem 12 (Zhang). There exists a unique finitely additive, convex-ranged²² probability measure P on Λ such that $A \succeq_{\ell} B \Leftrightarrow P(A) \ge P(B)$ for all $A, B \in \Lambda$ if and only if \succeq_{ℓ} satisfies:

- 1. $A \succeq_{\ell} \emptyset$ for any $A \in \Lambda$.
- 2. $X \succ_{\ell} \emptyset$.
- 3. \succeq_{ℓ} is a weak order.
- 4. If $A, B, C \in \Lambda$ and $A \cap C = B \cap C = \emptyset$, then $A \succ_{\ell} B$ if and only if $A \cup C \succ_{\ell} B \cup C$.
- 5. For any two uniform partitions $\{A_i\}_{i=1}^n$ and $\{B_i\}_{i=1}^n$ of S in Λ , $\bigcup_{i \in I} A_i \sim_{\ell} \bigcup_{i \in J} B_j$ if |I| = |J|.²³
- 6. (a) If $A \in \Lambda$ and $A \succ_{\ell} \emptyset$, there is a finite partition $\{A_1, \ldots, A_n\}$ of X in Λ such that: i. $A_i \subseteq A$ or $A_i \subseteq A^{\complement}$ for all A_i , and
 - *ii.* $A \succ_{\ell} A_i$ for all A_i .
 - (b) If $A, B, C \succ_{\ell} \emptyset$, $A \cup C = \emptyset$, and $B \succ_{\ell} A$, then there is a finite partition $\{C_1, \ldots, C_n\}$ of C in Λ such that $B \succ_{\ell} A \cup C_i$ for all C_i .
- 7. If $\{A_n\}$ is a decreasing sequence in Λ and $A^* \succ_{\ell} \bigcap_n A_n \succ_{\ell} A_*$ for some A^* and A_* in Λ , then there exists N such that $A^* \succ_{\ell} A_n \succ_{\ell} A_*$ for all $n \geq N$.

We can apply this theorem to Λ_A and the resulting quantitative probability has some nice properties in terms of representing \succeq on a restricted domain. We begin by proving a technical step.

Lemma 13. For all $A \subseteq X$, there exists $A_0, A_1 \subset A$ such that $A_0 \cap A_1 = \emptyset$, $A_0 \cup A_1 = A$, and $A_0 \cup B \sim A_1 \cup B$ for any $B \nsim A$.

Proof. Let $\overline{A} = \{x \in A : \{x\} \succeq A\}$ and $\underline{A} = \{x \in A : A \succ \{x\}\}$. Then, by an argument similar to the proof of Lemma 8, $\overline{A} \succ \underline{A}$. Fix some $\overline{x} \in \overline{A}$ and $\underline{x} \in \underline{A}$. Let $\overline{A}_{\alpha} = \alpha\{\overline{x}\} + (1 - \alpha)\overline{A}$ and $\underline{A}_{\alpha} = \alpha\{\underline{x}\} + (1 - \alpha)\underline{A}$. Since $\underline{A} \setminus \underline{A}_{\beta} \prec A$, disjoint set betweenness forces $\overline{A} \cup \underline{A}_{\beta} \succ A$. Fix $\beta \in [0, 1]$. By construction, $A \succ \underline{A}$. Then Lebesgue continuity implies there exists some $\alpha \in [0, 1]$ such that $\overline{A}_{\alpha(\beta)} \cup \underline{A}_{\beta} \sim A$. Moreover, this α is unique by disjoint set betweenness. Let the function $\alpha(\beta) : [0, 1] \rightarrow [0, 1]$ denote this assignment for each β , which is continuous by Lebesgue continuity. Thus $\overline{A}_{\alpha(\beta)} \cup \underline{A}_{\beta} \sim A$. By disjoint set betweenness, $A \setminus [\overline{A}_{\alpha(\beta)} \cup \underline{A}_{\beta}] \sim A$ as well.

Recall f(A) is a Lebesgue continuous utility representation of \succeq . Fix $B \prec A$; the argument for $B \succ A$ is parallel. Let $F(\beta) = f(\overline{A}_{\alpha(\beta)} \cup \underline{A}_{\beta} \cup B)$ and $G(\beta) = f(A \setminus [\overline{A}_{\alpha(\beta)}) \cup \underline{A}_{\beta}] \cup B)$. By construction, $G(0) = f(A \setminus [\overline{A}_{\alpha(\beta)}) \cup \underline{A}_{\beta}] \cup B$.

 $^{^{22}}$ A set function P on Λ is *convex-ranged* if for all $A \in \Lambda$ and $0 < \alpha < 1$, there exists some $B \subset A$ such that $P(B) = \alpha P(A)$. Notice this is much stronger than asserting that $P(\Lambda)$ is convex.

²³A partition $\{A_i\}_{i=1}^n$ is uniform if $A_i \sim_{\ell} A_j$ for all i, j.

 $f(A \cup B), \ G(1) = f(B), \ F(0) = f(B), \ \text{and} \ F(1) = f(A \cup B).$ Letting $H(\beta) = G(\beta) - F(\beta)$, we have $H(0) = f(A \cup B) - f(B) > 0 > f(B) - f(A \cup B) = H(1)$. As H is continuous in α , the Intermediate Value Theorem implies there exists β^* such that $H(\alpha^*) = 0$, i.e. $\overline{A}_{\alpha(\beta^*)} \cup \underline{A}_{\beta^*} \cup B \sim (A \setminus [\overline{A}_{\alpha(\beta^*)} \cup \underline{A}_{\beta^*}]) \cup B$. By balancedness, showing this indifference relation for a particular B proves it for all such B.

Lemma 14. Take any $A \succ B$ with $A \cap B = \emptyset$. There exist finitely additive, convex-ranged, tight²⁴ probability measures P_A and P_B on Λ_A and Λ_B such that:

- 1. $u(S \cup B) = P_A(S)$ is a utility representation of \succeq on $\{S \cup B : S \in \Lambda_A\}$; and
- 2. $u(A \cup T) = -P_B(T)$ is a utility representation of \succeq on $\{A \cup T : T \in \Lambda_B\}$.

Furthermore, P_A is robust to choice of $B \prec A$ and P_B is robust to choice of $A \succ B$.

Proof. Define the likelihood ordering \succeq_{ℓ} on Λ_A by $S_1 \succeq_{\ell} S_2$ if and only if $S_1 \cup B \succeq S_2 \cup B$ and similarly for Λ_B .

Step 1: Conditions 1, 2, 3, and 4. Conditions 1 and 2 are immediate consequences of disjoint set betweenness. Condition 3 follows since \succeq is a weak order. Condition 4 follows immediately from balancedness.

Step 2: If A_0 is a strict subset of A_1 , then $A_1 \succ_{\ell} A_0$. Since Λ_A is closed under disjoint unions, $A_1 \setminus A_0 \sim A$. We have $A \sim A_0 \succ B$, so disjoint set betweenness implies $A \succ A_0 \cup B$. Also by disjoint set betweenness, $A_1 \setminus A_0 \succ A_0 \cup B$ implies $A_1 \cup B = (A_1 \setminus A_0) \cup A_0 \cup B \succ A_0 \cup B$. By definition, this means $A_1 \succ_{\ell} A_0$.

Step 3: Suppose $A_1 \cap A_2 = A'_1 \cap A'_2 = \emptyset$, $A_1 \sim A_2$, and $A'_1 \sim A'_2$. If $A_1 \succ_{\ell} A'_1$, then $A_1 \cup A_2 \succ_{\ell} A'_1 \cup A'_2$. By divisibility, there exists disjoint B_1, B_2 such that $B_1 \cup B_2 = B$ and $A_1 \cup B_1 \sim A_1 \cup B_2$. By balancedness on B_1, B_2 , this implies $A_2 \cup B_1 \sim A_2 \cup B_2$. The definition of \succeq_{ℓ} implies $A_1 \cup B \sim A_2 \cup B$. Then balancedness on A_1, A_2 implies $A_1 \cup B_1 \sim A_2 \cup B_1$. Transitivity of \succeq on the previous indifference relations implies $A_1 \cup B_1 \sim A_2 \cup B_2$. Then disjoint set betweenness forces $A_1 \cup B_1 \cup A_2 \cup B_2 = A_1 \cup A_2 \cup B \sim A_1 \cup B_1$. A parallel argument establishes that $A'_1 \cup A'_2 \cup B \sim A'_1 \cup B_1$. Since $A_1 \succ_{\ell} A'_1$, we have $A_1 \cup B \succ A'_1 \succ B$. Then balancedness implies $A_1 \cup B_1 \succ A'_1 \cup B_1$. Thus $A_1 \cup A_2 \cup B \succ A'_1 \cup A'_2 \cup B$, i.e. $A_1 \cup A_2 \succ_{\ell} A'_1 \cup A'_2$.

Step 4: For any *n*, there exists a uniform partition $\{B_i\}_{i=1}^{2^n}$ of *B*. The proof is by induction. By divisibility, there exists a partition $\{B_1, B_2\}$ of cardinality 2. Now, suppose there exists a uniform partition $\{B_1, B_2\}$ of cardinality 2. Now, suppose there exists a uniform partition $\{B_1, B_2\}$ of $B_1, \ldots, B_{2^n}\}$. By applying divisibility to each B_k , we can produce a two-element partition $\{B_k^1, B_k^2\}$ of B_k such that $B_k^1 \cup A \sim B_k^2 \cup A$, i.e. $B_k^1 \sim_{\ell} B_k^2$. Step 3 implies $B_k^1 \sim_{\ell} B_m^1$ for all $k, m; B_k^1 \succ_{\ell} B_m^1$ would imply $B_k \succ_{\ell} B_m$, which would contradict $B_k \sim_{\ell} B_m$. Thus, $\{B_1^1, B_1^2, \ldots, B_{2^n}^1, B_{2^n}^2\}$ is a uniform partition of cardinality 2^{n+1} .

Step 5: Condition 5. Suppose $\{A_i\}_{i=1}^n$ and $\{A'_i\}_{j=1}^n$ are uniform partitions of A. By the previous step, there exists a sequence of disjoint subsets B_1, \ldots, B_n of B such that $B_i \sim B$ and $B_i \sim_{\ell} B_j$ for all i, j. Let $B_0 = \bigcup_{i=1}^n B_i$.

We first prove that $A_i \sim A'_j$ for all i, j. To the contrary, assume without loss of generality that $A_1 \succ_{\ell} A'_1$. Then $A_i \sim_{\ell} A_1 \succ_{\ell} A'_1 \sim_{\ell} A'_j$. By two applications of balancedness, $A_1 \cup B_1 \sim A_i \cup B_1 \sim A_i \cup B_i$ and similarly $A'_1 \cup B_1 \sim A'_j \cup B_j$. By construction, $A \cup B_0 = \bigcup_{i=1}^n A_i \cup B_i = \bigcup_{j=1}^n A'_j \cup B_j$. But disjoint set betweenness iteratively applied to $A_i \cup B_i \succ A'_i \cup B_i$ implies $\bigcup_{i=1}^n A_i \cup B_i \succ \bigcup_{j=1}^n A'_j \cup B_j$, a contradiction.

Now consider $\bigcup_{i=1}^{m} A_i$ and $\bigcup_{j=1}^{m} A'_j$ for some $m \leq n$. We have just shown that $A_i \sim_{\ell} A'_i$ for all *i*. Then disjoint set betweenness iteratively applied to $A_i \cup B_i \sim A'_i \cup B_i$ implies $\bigcup_{i=1}^{m} (A_i \cup B_i) \sim \bigcup_{j=1}^{m} (A'_j \cup B_j)$, i.e. $(\bigcup_{i=1}^{m} A_i) \cup (\bigcup_{i=1}^{m} A'_j) \cup (\bigcup_{j=1}^{m} B_j)$. Then balancedness implies $(\bigcup_{i=1}^{m} A_i) \cup B \sim (\bigcup_{j=1}^{m} A'_j) \cup B$. By definition, this means $\bigcup_{i=1}^{m} A_i \sim_{\ell} \bigcup_{j=1}^{m} A'_j$.

²⁴A set function P is tight if for every set $A \in \Lambda$, $P(A) = \sup\{P(K) : K \in \Lambda, K \text{ compact}, K \subseteq A\}$.

Step 6: Condition 6. We first prove part (a). Fix $A^0 \subseteq \Lambda_A$ such that $A_0 \neq \emptyset$ and let $A^1 = A \setminus A^0$. It suffices to show that there exists a finite partition $\{A_1, \ldots, A_n\}$ of A such that $A_i \subseteq A^0$ or $A_i \subseteq A^1$ and that $A_i \succ_{\ell} A^0, A^1$. If $A_1 \sim_{\ell} A_0$, we can use divisibility to split both sets and we are done. So, without loss of generality, suppose $A^1 \succ_{\ell} A^0$. By Step 4, for any m, we can find a uniform partition $\mathcal{A}^{(m)} = \{A_1, \ldots, A_{2m}\}$ of A^1 . By additivity of measure, there must exist $A_i^1 \in \mathcal{A}^{(m)}$ with $\lambda(A_i) < 2^{-m}$. Then there exists a sequence of uniform partitions $\mathcal{A}^{(1)}, \mathcal{A}^{(2)} \ldots$ of A^1 such that $\lambda(A_1^{(m)}) \to 0$. Suppose $A_1^{(m)} \succeq_{\ell} A^0$ for all m. Then, by Condition 4, $A_1^{(m)} \cup (A^1 \setminus A_1^{(m)}) \succeq_{\ell} A^0 \cup (A \setminus A_1^{(m)})$. But the left side of the relation is exactly A^1 and the right hand side converges in λ to A. Then Lebesgue continuity implies $A^1 \succeq_{\ell} A$, a contradiction of Step 2. So, there exists a uniform partition $\{A_1^1, \ldots, A_m^1\}$ such that $A^0 \succ A_i^1$ for all i. The same argument, applied to $A_1^{(m)}$ and A^0 , provides a uniform partition $\{A_1^0, \ldots, A_n^0\}$ of A^0 such that $A_1^{(m)} \succ_{\ell} A_1^0$, a fortiori that $A^1 \succ_{\ell} A_1^0$. Then the union of the two partitions is the required partition of ΔX .

Part (b) similarly follows from Lebesgue continuity and Step 4.

Step 7: Condition 7. This following immediately from Lebesgue continuity.

Later in the proof, we require a weighting on all sets indifferent to A. This step is not immediate because the family of all sets indifferent to A does not contain a superset to use as X. We now extend P_A to the entire level set.

The first application of Lemma 14 is in proving an intermediate preference result. In words, if A is preferred to B, we can "calibrate" preference by taking subsets of either the better set A or the worse set B.

Lemma 15. Suppose $A \succeq B$ and $A \cap B = \emptyset$. If $A \succeq C \succeq A \cup B$, then there exists $B' \in \Lambda_B$ such that $A \cup B' \sim C$. If $A \cup B \succeq C \succeq B$, then there exists $A' \in \Lambda_A$ such that $A' \cup B \sim C$.

Proof. If $A \sim B$, the result follows immediately from disjoint set-betweenness using A, B as A', B'. So take $A \succ B$ and assume the first case, $A \succ C \succ A \cup B$. Let P_B refer to the measure provided by Lemma 14 and the corresponding representation u. Take any continuous utility representation f of \succeq over all \mathcal{Z} , normalized so f = u on $\{A \cup T : T \in \Lambda_B\}$. Then f(A) = 0; $f(A \cup B) = -1$; and $-1 \le f(C) \le 0$. Recall that Lemma 14 also states that $u(A \cup T) = -P_B(T)$ for any $T \in \Lambda_B$. Since P_B is convex-valued, there exists some $B' \in \Lambda_B$ such that $P_B(B') = -f(C)$. Then $f(A \cup B') = u(A \cup B') = -P_B = f(C)$, which proves the first statement of the Lemma. The proof of the second statement is symmetric.

Lemma 16. There exists a finitely additive set function that extends P_A to the level set of A, which is unique up to a scale transformation.

Proof. Fix $A \in \mathcal{Z}$. If $A \sim \Delta X$, then the extension is already provided by $P_{\Delta X}$. So, we may assume that $A \nsim \Delta X$. We assume $A \succ \Delta X$, the case $A \prec \Delta X$ is entirely analogous. Suppose $B \sim A$. Case 1: There exists $b \in B$ such that $\{b\} \sim B$. Then, since $A \succ \Delta X$, by downward Hausdorff continuity we can assume that there exists a set $B' \subseteq B$ containing b such that $\lambda(B') > 0$. Moreover, by π -continuity, since either $b \in A$ or $b \notin A$, we can assume without loss of generality that $B' \subseteq A$ or $B' \cap A = \emptyset$. If it is the former, set $P_A(B) = \frac{P_A(B')}{P_{B(B')}}$. If it is the latter, then $B' \cup A \sim A$ by disjoint set betweenness. Then set $P_A(B) = \frac{P_B(B')}{P_{A\cup B'}(B')P_{A\cup B'}(A)}$. Case 2: There exists some $b \notin B$ such that $\{b\} \sim B$. By downward Hausdorff continuity, we can assume there exists a set B' disjoint from B and containing b such that $B' \sim B$. Then find $P_A(B' \cup B)$ by Case 1. Set $P_A(B) = P_{B' \cup B}(B)P_A(B' \cup B)$. To observe finite additivity, observe that if B_1 and B_2 are disjoint and $B_1 \sim B_2 \sim A$, then $B_1 \cup B_2 \sim A$ by disjoint set betweenness and additivity of P_A then follows from the additivity of $P_{B_1 \cup B_2}$.

In Lemma 14, we constructed a measure that represents preference between certain unions of a fixed set S and the members of Λ_A . We now construct a signed measure whose sign indicates preference relative to a fixed set A.

Lemma 17. If $A \in \mathbb{Z}$, then there exists a nonatomic, finitely additive, tight signed measure ν_A on \mathbb{Z} such that: $\nu_A(S) \geq 0$ if and only if $S \succeq A$.

Proof. For a fixed A, let $\overline{L} = \{l \in \Delta X : l \prec A\}$ and $\mathcal{L} = \{L \in \mathcal{Z} : L \subseteq \overline{L}\}$. Similarly, let $\overline{U} = \{l \in \Delta X : l \succ A\}$ and $\mathcal{U} = \{U \in \mathcal{Z} : U \subseteq \overline{U}\}$. Then \mathcal{L} and \mathcal{U} are respectively σ -algebras of \overline{L} and \overline{U} , after appending the empty set. Either $A \succeq \Delta X$ or $A \preceq \Delta X$. We will assume the former; the argument for the second case is similar.

Fix any $U \in \mathcal{U}$. By our intermediate calibration result, Lemma 15, there exists some $L \in \Lambda_{\bar{L}}$ such that $U \cup L \sim A$. Set $\nu(U) = P_{\bar{L}}(L)$, where $P_{\bar{L}}$ is produced by Lemma 14. By the construction of $P_{\bar{L}}$, $\nu(U)$ is robust to our choice of L, hence uniquely defined.

Select $U_1, U_2 \in \mathbb{Z}$ with $U_1 \cap U_2 = \emptyset$. There exists some $L_{12} \in \Lambda_{\bar{L}}$ such that $L_{12} \cup U_1 \cup U_2 \sim A$, by Lemma 15. Without loss of generality, assume $U_1 \succeq U_2$. By disjoint set betweenness, $L_{12} \cup U_1 \succeq A$. Then we can apply Lemma 15 again to L_{12} to find $L_1 \subset L_{12}$ with $L_1 \cup U_2 \sim A$. Set $L_2 = L_{12} \setminus L_1$. Another application of disjoint set betweenness forces $U_2 \cup L_2 \sim A$, since $L_1 \cup L_2 \cup U_1 \cup U_2 = (L_1 \cup U_1) \cup (L_2 \cup U_2) \sim A$ and $L_1 \cup U_1 \sim A$. Therefore, ν inherits disjoint additivity from $P_{\bar{L}}$. We now have a finitely additive signed measure ν on \mathcal{U} .

Now take any $L \in \mathcal{L}$. If there exists some $U \in \mathcal{U}$ such that $L \cup U \sim A$, let $\nu(L) = -\nu(U)$. On the other hand, if there is no $U \in \mathcal{U}$ with $L \cup U \sim A$, we can use Lemma 15 to produce a subset $L' \in \Lambda_L$ such that there exists $U' \in \mathcal{U}$ with $L' \cup U' \sim A$. Let P_L refer to the measure on Λ_L produced by Lemma 14. Set $\nu(L) = \frac{\nu(U')}{P_L(L')}$. The measure ν inherits disjoint additivity on \mathcal{L} from \mathcal{U} by its construction. This extends ν to \mathcal{L} .

We move to any arbitrary $S \in \mathbb{Z}$. If $s \sim A$ for all $s \in S$, set $\nu(S) = 0$. Otherwise, we can express this set as $S = L \cup U$ for some $L \in \mathcal{L}$ and $U \in \mathcal{U}$. Set $\nu(A) = \nu(L) + \nu(U)$. Additivity is immediately inherited from \mathcal{L} and \mathcal{U} . We have now extended ν to all of \mathbb{Z} .

Verifying the representation claim, suppose $S \succeq A$. Then $(S \cap \overline{L}) \cup (S \cap \overline{U}) \succeq A$. By Lemma 15, we can find a subset $U' \in \Lambda_{S \cap \overline{U}}$ with $(S \cap \overline{L}) \cup U' \sim A$. Since $U' \subseteq S \cap \overline{U}$, $\nu(U') \leq \nu(S \cap \overline{U})$. Recalling the construction, $\nu(S) = \nu(S \cap \overline{L}) + \nu(S \cap \overline{U}) = \nu(S \cap \overline{U}) - \nu(U') \geq 0$. Similar arguments establish that $\nu(S) > 0$ only if $S \succ A$.

The measure ν_A is convex-ranged, hence nonatomic. Tightness is inherited from P_A by construction. \Box

Up to this point, we have only considered finitely additive measures, sometimes called charges. We now show that they are also countably additive, which is intuitively a consequence of Lebesgue continuity.

Lemma 18. The measure ν_A is countably additive.

Proof. Each ν_A is finite and ΔX is a Hausdorff space. Since ν_A is tight by Lebesgue continuity, (Aliprantis and Border 1999, Theorem 10.4) implies ν_A is countably additive.

Now consider $ca(\mathcal{Z})$, the set of all countably additive finite (signed) measures on \mathcal{Z} , which is a topological vector space under the topology of weak convergence. The representation features of ν_A are robust to positive scalar transformations, i.e. $\alpha\nu_A$ has the same properties whenever $\alpha > 0$. The measures constructed in Lemma 17 live in $ca(\mathcal{Z})$. The next result shows that these measures can be spanned by two elements of $ca(\mathcal{Z})$. This base will become the critical part of the representation.



Figure 3: Lemma 19

This three-dimensional space is supposed to capture the infinite-dimensional space $ca(\mathcal{Z})$. Drawn with heavier lines, the plane cutting through the origin is the span of ν and μ . The measures ν_A , drawn as the dotted curve, live inside that span. A particular ν_A is labelled inside the curve.

Lemma 19. The family $\{\nu_A : A \in \mathcal{Z}\}$ is spanned by two measures ν, μ , with μ a probability measure.

Proof. Take any A, B, C which are not indifferent to each other. We lose no generality by ordering them $A \succ B \succ C$. All the measures ν_A are nonatomic. We can invoke the Lyapunov Convexity Theorem: the range of the vector-valued measure $[\nu_A, \nu_B, \nu_C]$,

$$[\nu_A, \nu_B, \nu_C](\mathcal{Z}) = \{(\nu_A(S), \nu_B(S), \nu_C(S)) \in \mathbb{R}^3 : S \in \mathcal{Z}\},\$$

is convex.²⁵ Take any S with $\nu_A(S) = 0$. By construction, $\nu_B(S) > 0$ and $\nu_C(S) > 0$. By using Lemma 15, we can find $S^* \in \Lambda_S$ with $S^* \cup L_1 \sim B$ and $S^* \cup L_2 \sim C$ for some L_1, L_2 disjoint from S^* . Recalling the representation condition in Lemma 17, $\nu_B(S^*) + \nu_B(L_1) = \nu_B(S^* \cup L_1) = 0$; similarly $\nu_C(S^*) + \nu_C(L_1) = 0$. Now take any other T with $\nu_A(T) = 0$, and assume without loss of generality that $T \cap S^* = 0$. Then, suppose $\nu_B(T) \geq \nu_B(S^*)$. Then $\nu_C(T) \geq \nu_C(S^*)$, by disjoint set betweenness and the representation condition applied again. The same arguments hold for the strict inequality as well. Therefore ν_B and ν_C induce the same ordering on $\{S: S \sim A\}$. Applying Lemma 16 and the uniqueness claim of Theorem 12, this ordering completely determines ν_B and ν_C on this restricted domain up to a scale transformation, i.e. $\nu_B(S) = c\nu_C(S)$ for a positive constant c, across any S with $\nu_A(S) = 0$. Then $\{x \in \mathbb{R}^3 : x_1 = 0\} \cap [\nu_A, \nu_B, \nu_C](\mathcal{Z})$ is contained in the ray $\{(0, ct, t) : t \geq 0\}$.

Since ν_A has strictly positive components (namely the strict upper contour set of A), the set

$$\{x \in [\nu_A, \nu_B, \nu_C](\mathcal{Z}) : x_1 > 0\}$$

is nonempty. Therefore, we can select vectors $x^0 \in [\nu_A, \nu_B, \nu_C](\mathcal{Z})$ such that $x_1^0 = 0$ and $x^1 \in [\nu_A, \nu_B, \nu_C](\mathcal{Z})$ such that $x_1^1 > 0$. We show that the two vectors x^0, x^1 together span $[\nu_A, \nu_B, \nu_C](\mathcal{Z})$. Let $x^* \in [\nu_A, \nu_B, \nu_C](\mathcal{Z})$. We proceed by cases.

²⁵We thank Yossi Feinberg for suggesting the Lyapunov Convexity Theorem, which simplified an earlier proof.

Case 1: Suppose $x_1^* = 0$. Since $\{x \in \mathbb{R}^3 : x_1 = 0\} \cap [\nu_A, \nu_B, \nu_C](\mathcal{Z})$ is a ray, $x^* = cx^0$.

Case 2: Suppose $x_1^* < 0$. Since $x_1^1 > 0$ is nonempty, we can find a convex combination $x' = \alpha x^* + (1-\alpha)x^1$ with $x_1' = 0$. The vector x' is in the range of the vector-valued measure $[\nu_A, \nu_B, \nu_C]$, because it is a convex combination of two elements. But, applying Case 1 to x', we conclude x' is spanned by x^0 . x^* is spanned by x' and x^1 , while x' is a multiple of x^0 . So x^* is spanned by x^0 and x^1 .

Case 3: Suppose $x_1^* > 0$. Apply Case 2 to $-x_1^*$.

Hence, there exist constants α, β such that $\nu_A = \alpha \nu_B + \beta \nu_C$. This suffices to show the entire space $\{\nu_A : A \in \mathcal{Z}\}$ can be spanned by any two of its measures, since the selection of A, B, C in the proof was arbitrary.

We finally show this span contains a strictly positive measure. Rescale each ν_A so $||\nu_A|| = 1$. The space of probability measures $\Delta(\mathcal{Z})$ is a closed subset of $ca(\mathcal{Z})$, while ν_A approaches $\Delta(\mathcal{Z})$ as A approaches the \succeq -minimal element a_* in the order topology of \succeq , which is coarser than the Euclidean topology by continuity. Since any subspace of $ca(\mathcal{Z})$ is closed, the span contains a probability measure, which we can use as μ . \Box

Lemma 20. The set $\{\alpha\nu_A : A \in \mathbb{Z}, \alpha > 0\}$ is a convex positive cone.

Proof. Take ν_A, ν_B . If $A \sim B$, then $\nu_A = \alpha \nu_B$ for some constant $\alpha > 0$ and any convex combination of ν_A and ν_B is immediately in $\{\alpha \nu_A : A \in \mathcal{Z}, \alpha > 0\}$. So assume that $A \succ B$. Then ν_A and ν_B are linearly independent and can serve as a basis for span (ν, μ) . Let $\nu_\alpha = \alpha \nu(A) + (1 - \alpha)\nu_B$.

There exists some C such that $\nu_{\alpha}(C) = 0$. Since $\nu_A \ge 0, \nu_B(C) > 0$ whenever $C \succeq A$ and $\nu_B(C) \le 0, \nu_A(C) < 0$ whenever $C \preceq B$, it must be the case that $A \succ C \succ B$. Then ν_B, ν_C are linearly independent and can serve as a basis for span (ν, μ) . Pick $\gamma < \frac{\nu_C(A)}{\nu_B(A)}$. Notice $\gamma > 0$ since $\nu_C(A), \nu_B(A) > 0$. Since ν_B, ν_C are linearly independent, $\gamma \nu_C, \nu_B$ are also linearly independent and can serve as a basis for span (ν, μ) . Let (q, r) denote the coordinates for ν_A with respect to this basis.

Since $\nu_A(A) = 0$ we have $q(\nu_B(A)) + r(\gamma\nu_C(A)) = \nu_A(A) = 0$. By our selection of γ , we have q < 0 < r. Let $c = 1/r\gamma > 0$ and $d = -q/r\gamma > 0$. Then $\nu_C = c\nu_A + d\nu_B$. Let $\beta = c/(c+d) > 0$. Simple algebra verifies that $\nu_\alpha = \beta\nu_C$.

Lemma 21. There exists some ν^* in $span(\nu, \mu)$ such that $\frac{\nu^*(A)}{\mu(A)}$ is a utility representation.

Proof. For any $\gamma \in \operatorname{span}(\mu, \nu)$, let $\alpha(\gamma), \beta(\gamma)$ solve $\alpha(\gamma)\nu + \beta(\gamma)\mu = \gamma$. Take any $B \succeq C$. If $B \sim C$, then $\nu_B = \nu_C$. If $B \succ C$, there exists some D with $B \succ D \succ C$, implying $\nu_B(D) > 0$ and $\nu_C(D) < 0$. Since $\nu_B(A) > 0$ and $\nu_C(A) > 0$, there exists no $\alpha > 0$ such that $\nu_B = -\alpha\nu_C$. In either case, it is impossible for $\nu_B = -\alpha\nu_C$ for any $\alpha > 0$. As B, C are arbitrary and $\{\gamma\nu_S : S \in \mathcal{Z}, \gamma > 0\}$ is a convex positive cone by Lemma 20, there exists some half space H in $ca(\mathcal{Z})$ such that $\{\nu_S : S \in \mathcal{Z}\} \subseteq H$. Let $H^* = H \cap \operatorname{span}(\mu, \nu)$. This H^* is a two-dimensional half space of $\operatorname{span}(\mu, \nu)$, so H^* is defined by $\{\gamma \in \operatorname{span}(\mu, \nu) : a\alpha(\gamma) + b\beta(\gamma) + c = 0\}$ for some constants a, b, c.

Furthermore, for any $S \in \mathbb{Z}$, there exists $T \succ S$, for which $\nu_S(T) > 0$ and $\mu(T) \ge 0$. Therefore, there exists no $\alpha > 0$ or ν_S such that $\nu_S = -\alpha\mu$. Therefore, we may proceed without loss of generality by assuming $a\alpha(\mu) + b\beta(\mu) + c = 0$. Pick $\nu^* \in \operatorname{span}(\mu, \nu)$ such that $a\alpha(\nu^*) + b\beta(\nu^*) > 0$. Obviously, μ and ν^* are linearly independent. Therefore, we can find $\alpha^*(A), \beta^*(A)$ such that $\alpha^*(A)\nu^* + \beta^*(A)\mu = \nu_A$. By the selection of ν^* , we must have $\alpha^*(\nu_A) > 0$ to satisfy the inequality $\alpha(\nu_A) + \beta(\nu_B) + c \ge 0$.

Let $\alpha(A) = \alpha^*(\nu_A)$ and $\beta(A) = -\beta^*(\nu_A)$. By construction,

$$\alpha(A)\nu^*(A) - \beta(A)\mu(A) = 0 = \nu_A(A).$$



Figure 4: Constructions in the proof of Lemma 21

This two-dimensional figure shows the span of ν and μ , laying the plane in Figure 3 flat against the page. The line ∂H is the border that defines the half space H in which all of the ν_A 's live; ν is normal to that border. The coefficients $\alpha(A)$ and $\beta(A)$ on ν and μ are shown for a particular set A; the coefficients are defined by $\nu_A = \alpha(A)\mu - \beta(A)\nu$.

Then

$$\frac{\beta(A)}{\alpha(A)} = \frac{\nu^*(A)}{\mu(A)},$$

so it suffices to show that the left hand side is a representation for \succeq . Take $A \succeq B$. By construction,

$$\nu_B(A) \ge 0 = \nu_A(A).$$

Also, our selection of $\alpha(A)$ and $\beta(A)$ implies

$$\nu(A) - \frac{\beta(A)}{\alpha(A)}\mu(A) = \frac{1}{\alpha(A)} [\alpha(A)\nu^*(A) - \beta(A)\mu(A)]$$
$$= \frac{\nu_B(A)}{\alpha(A)}$$
$$\geq 0,$$

the last inequality following from $\alpha(A) > 0$. Similarly,

$$\nu^*(A) - \frac{\beta(B)}{\alpha(B)}\mu(A) \le 0.$$

Together these two inequalities imply

$$\frac{\beta(A)}{\alpha(A)} \ge \frac{\beta(B)}{\alpha(B)}.$$

The argument when $A \succ B$ is identical, replacing weak inequalities with strict inequalities.

Proof of Theorem. Invoking the Radon–Nikodym Theorem, we can rewrite $\nu^*(A)$ of Lemma 21 as $\int_A u \, d\mu$, where u is the Radon–Nikodym derivative of ν with respect to μ . The absolute continuity condition on these measures holds because $\mu(A) > 0$ for any $A \in \mathbb{Z}$ by disjoint set betweenness. The continuity of u is guaranteed by the downward Hausdorff continuity axiom and Lemma 7. the full support of μ follows from the strict part of disjoint set betweenness. This proves the sufficiency of the axioms for the representation on the regular sets. The representation for singletons follows immediately from Lemma 7.

A.2 Proof of Theorem 6

We begin by introducing some useful notation. Let $A + y = \{a + y : a \in A\}$ for any signed measure y on ΔX . Then A + y is a translation of A within ΔX if $y(\Delta X) = 0$ and each $a + y \in A + y$ is positive, because any mass taken away from sets by y will be assigned to other sets, preserving total mass at unity. We first prove that, in our model, singleton-set independence implies a form of translation invariance.

Axiom 8 (Translation independence). For any regular $A \subseteq \Delta X$, $a \in \Delta X$, and signed measure y on ΔX such that $\{a + y\}, A + y \in \mathcal{K}^*, A \sim \{a\}$ if and only if $\{a + y\} \sim A + y$.

Lemma 22. If \succeq is singleton-set independent, then \succeq is translation independent.

Proof. Take $a \sim A$ and y with $a + y, A + y \in \mathcal{K}^*(\Delta X)$. Simple algebra verifies

$$(1-\alpha)A + \alpha(a+y) = \alpha[(1-\alpha)a + \alpha(a+y)] + (1-\alpha)[A+\alpha y].$$

Then reflexivity implies

$$(1-\alpha)A + \alpha(a+y) \sim \alpha[(1-\alpha)a + \alpha(a+y)] + (1-\alpha)[A+\alpha y].$$

By singleton-set independence,

$$(1-\alpha)A + \alpha(a+y) \sim (1-\alpha)a + \alpha(a+y).$$

These two indifference relations and another application of singleton-set independence force

$$(1 - \alpha)a + \alpha(a + y) \sim A + \alpha y.$$

Then, taking $\alpha \to 1$ and invoking Lebesgue continuity obtains $a + y \sim A + y$. The other direction is similar, expressing A = (A + y) + (-y).

Fix the representation (ϕ, u, μ) delivered by the linearity axiom. The space of probability measures spans the space of all signed measures. So there is a unique extension of $\phi \circ u$ to the space of all signed measures since ϕ is linear. We work with this extension and denote $U(y) = \phi(u(y))$. Furthermore, f refers to the Radon–Nikodym derivative $\frac{d\mu}{d\lambda}$ where λ is the (|X| - 1)-dimensional Lebesgue probability measure on ΔX . We first show this derivative exists.

Lemma 23. The Radon-Nikodym derivative $f = \frac{d\mu}{d\lambda}$ exists.

Proof. Lemma 10 of Appendix A.1 shows that the extended preference is continuous with respect to the π metric generated by λ on the measure σ -algebra $\mathcal{B}(\Delta X)/\mathcal{N}$. Then μ must be absolutely continuous with respect to λ . If not, there exists a set A with $\lambda(A) > 0$ and a disjoint set $B \succ A$ such that $A \cup B \prec B$. But $\lambda([A \cup B]\Delta B) = 0$, and this contradicts continuity of the extended preference.

Lemma 24.

 $U(\alpha A + (1 - \alpha)x) = \alpha U(A) + (1 - \alpha)U(\{x\})$

and

$$U(A+y) = U(A) + U(y).$$

Proof. The ambiguity-free equivalent of a set A is identified as any point x with $U(\{x\}) = U(A)$. Since ϕ is strictly increasing and linear, it has an inverse, so we can equivalently write $u(x) = \phi^{-1}(U(\{x\}))$. Singleton-set independence implies

$$\phi^{-1}[U(\alpha A + (1-\alpha)x)] = \alpha \phi^{-1}[U(A)] + (1-\alpha)\phi^{-1}[U(\{x\})].$$

Applying ϕ to both sides of the equation yields

$$U(\alpha A + (1 - \alpha)x) = \phi \left(\alpha \phi^{-1}[U(A)] + (1 - \alpha)\phi^{-1}[U(\{x\})]\right)$$

= $\alpha U(A) + (1 - \alpha)U(\{x\}),$

where the last step follows from linearity, which allows us to distribute ϕ across the mixture. By Lemma 22,

$$\phi^{-1}U(A+y) = \phi^{-1}U(A) + \phi^{-1}U(y).$$

Similarly, this implies U(A + y) = U(A) + U(y).

Lemma 25.

$$\frac{\mu(A)}{\mu(A+y)} = \frac{\mu(B)}{\mu(B+y)}.$$

Proof. Take any ambiguous sets A and B with $A \succ B$ and any signed measure y with $U(y) \neq 0$. We can assume they are disjoint without loss of generality, by splitting their union into $A \setminus B$, $B \setminus A$, and $A \cap B$. Then

$$U(A \cup B) = \frac{\mu(A)}{\mu(A \cup B)}U(A) + \frac{\mu(B)}{\mu(A \cup B)}U(B).$$

By the second half of Lemma 24:

$$\begin{array}{lll} U(A \cup B) + U(y) &=& U([A \cup B] + y) \\ &=& U([A + y] \cup [B + y]) \\ &=& \frac{\mu(A + y)}{\mu([A \cup B] + y)} U(A + y) + \frac{\mu(B + y)}{\mu([A \cup B] + y)} U(B + y) \\ &=& \left[\frac{\mu(A + y)}{\mu([A \cup B] + y)} U(A) + \frac{\mu(B + y)}{\mu([A \cup B] + y)} U(B) \right] + U(y) \end{array}$$

Then

$$U(A \cup B) = \frac{\mu(A+y)}{\mu([A \cup B]+y)}U(A) + \frac{\mu(B+y)}{\mu([A \cup B]+y)}U(B)$$

These two equations imply that the weighting on U(A) in both convex combinations must be equal:

$$\frac{\mu(A)}{\mu(A\cup B)} = \frac{\mu(A+y)}{\mu([A\cup B]+y)}.$$

Rearranging terms,

$$\frac{\mu(A)}{\mu(A+y)} = \frac{\mu(A \cup B)}{\mu([A \cup B]+y)}$$

Symmetrically,

$$\frac{\mu(B)}{\mu(B+y)} = \frac{\mu(A \cup B)}{\mu([A \cup B]+y)}.$$

Therefore,

$$\frac{\mu(A)}{\mu(A+y)} = \frac{\mu(B)}{\mu(B+y)}.$$

Lemma 26.

$$\frac{\mu(A)}{\mu(\alpha A + (1 - \alpha)x)} = \frac{\mu(B)}{\mu(\alpha B + (1 - \alpha)x)}.$$

Proof. Again take A, B disjoint. The first half of Lemma 24 implies:

$$\begin{aligned} &\alpha U(A\cup B) + (1-\alpha)U(x) \\ &= U(\alpha[A\cup B] + (1-\alpha)x) \\ &= U([\alpha A + (1-\alpha)x] \cup [\alpha B + (1-\alpha)x]) \\ &= \frac{\mu(\alpha A + (1-\alpha)x)}{\mu(\alpha[A\cup B] + (1-\alpha)x)}U(\alpha A + (1-\alpha)x) + \frac{\mu(\alpha B + (1-\alpha)x)}{\mu(\alpha[A\cup B] + (1-\alpha)x)}U(\alpha B + (1-\alpha)x) \\ &= \alpha \left[\frac{\mu(\alpha A + (1-\alpha)x)}{\mu(\alpha[A\cup B] + (1-\alpha)x)}U(A) + \frac{\mu(\alpha B + (1-\alpha)x)}{\mu(\alpha[A\cup B] + (1-\alpha)x)}U(B)\right] + (1-\alpha)U(x). \end{aligned}$$

Then an argument similar to the end of the proof of Lemma 25 delivers the result.

Lemma 27. If f is continuous at any point a, it is continuous everywhere.

Proof. Suppose f is continuous at a. Pick a sequence $A_n \to \{a\}$. By Lemma 7 in Appendix A.1, $\frac{\mu(A_n)}{\lambda(A_n)} \to f(a)$. Any $b \in \Delta X$ can be expressed as the translation $x + y_b$ for some signed measure y_b and some $x \in \operatorname{int}(\Delta X)$. Letting $B_n = A_n + y_b$, we show $\frac{\mu(B_n)}{\lambda(B_n)}$ converges for any b. The numerator is a fixed multiple of $\mu(A_n)$ because the ratio $\frac{\mu(B_n)}{\mu(A_n)} = \frac{\mu(A_n + y_b)}{\mu(A_n)}$ is constant by Lemma 25; denote this constant by β . The denominator $\lambda(B_n) = \lambda(A_n + y) = \lambda(A_n)$, by the translation invariance of Lebesgue measure on ΔX . So $\frac{\mu(B_n)}{\lambda(B_n)} \to \beta\left(\lim \frac{\mu(A_n)}{\lambda(A_n)}\right)$. Then, we can consider the function $f(b) = \lim \frac{\mu(B_n)}{\lambda(B_n)}$ and this limit is robust for any sequence $B_n \to \{b\}$. Now take a sequence of measurable partitions Π_n of ΔX such that for any $b \in \Delta X$, there exists some $B_n \in \Pi_n$ with $B_n \to \{b\}$; for example, the partitions defined by finer grids in the Euclidean metric. Consider the simple functions defined by $f_n(x) = \frac{\mu(B_n)}{\lambda(B_n)}$ for all $x \in B_n \in \Pi_n$. The definition of the Lebesgue integral implies $f' = \lim f_n$ is a version of the Radon–Nikodym derivative. f' is continuous by construction in view of Lemma 7. So, if f is continuous at a we may take without loss of generality that f is continuous everywhere, as there is a version f' = f almost everywhere with f' continuous everywhere.

Lemma 28. f is continuous.

Proof. We begin by demonstrating Lemma 25 implies that almost surely either $f(a) \ge f(a+y)$ or $f(a) \ge f(a-y)$ for all directions y. Suppose not. Then there is a set A and a direction y with f(a+y) > f(a) for $a \in A$ μ -almost surely and a set B with f(b+y) < f(b) for $b \in B$ μ -almost surely with $\mu(A), \mu(B) > 0$. By translation invariance of Lebesgue measure, $\mu(A+y) = \int_{A+y} f(x) d\lambda = \int_A f(x+y) d\lambda > \int_A f(x) d\lambda = \mu(A)$. Similarly, $\mu(B+y) < \mu(B)$. Then

$$\frac{\mu(A)}{\mu(A+y)} < 1 < \frac{\mu(B)}{\mu(B+y)},$$

which contradicts Lemma 25.

This shows that, almost surely, f is monotone in all directions y of arbitrary length. Then, f is differentiable almost everywhere, hence continuous almost everywhere. Since this implies there exists at least one point where f is continuous, Lemma 27 implies f is continuous everywhere.

Proof of Theorem. Necessity is omitted and all functional equations are almost sure. Since f is continuous by Lemma 28, we can invoke Lemma 7 from Appendix A.1. It implies $f(a) = \lim \frac{\mu(A_n)}{\lambda(A_n)}$. Taking the limit of a sequence $A_n \to \{a\}$ and using Lemma 26:

$$\frac{f(a)}{f(\alpha a + (1 - \alpha)x)} = \frac{f(b)}{f(\alpha b + (1 - \alpha)x)}.$$

Lemma 25 similarly implies

$$\frac{f(a)}{f(a+y)} = \frac{f(b)}{f(b+y)}$$

Translating both a and a + y by y yields

$$\frac{f(a)}{f(a+y)} = \frac{f(a+y)}{f(a+2y)}.$$

Taking the convex combinations of $\frac{1}{2}a + \frac{1}{2}(a+2y) = a+y$ and $\frac{1}{2}(a+y) + \frac{1}{2}(a+2y) = a + \frac{3}{2}y$ yields

$$\frac{f(a)}{f(a+y)} = \frac{f(a+y)}{f(a+\frac{3}{2}y)}$$

Translating a + y and $a + \frac{3}{2}y$ by $\frac{1}{2}y$ yields

$$\frac{f(a+y)}{f(a+\frac{3}{2}y)} = \frac{f(a+\frac{3}{2}y)}{f(a+2y)}.$$

Combining these equalities:

$$\frac{f(a)}{f(a+y)} = \frac{f(a+y)}{f(a+2y)}$$

$$= \frac{f(a+y)}{f(a+\frac{3}{2}y)} \cdot \frac{f(a+\frac{3}{2}y)}{f(a+2y)}$$

$$= \left[\frac{f(a+y)}{f(a+\frac{3}{2}y)}\right]^2$$

$$= \left[\frac{f(a)}{f(a+y)}\right]^2$$

This forces f(a) = f(a + y). Since this is for arbitrary a, y and any point b can be expressed as a + y for some y, we conclude f(a) = f(b) for all a, b. Then the Radon–Nikodym derivative $\frac{d\mu}{d\lambda} = f$ is constant, so μ differs from λ only by a constant scaling. As μ and λ are both probability measures, this suffices to show $\mu = \lambda$.

References

AHN, D. S. (2003): "Hierarchies of Ambiguous Beliefs," Working paper, University of California, Berkeley.

ALIPRANTIS, C. D., AND K. C. BORDER (1999): Infinite Dimensional Analysis. Springer, New York, 2nd edn.

- ANSCOMBE, F. J., AND R. J. AUMANN (1963): "A Definition of Subjective Probability," Annals of Mathematical Statistics, 34, 199–205.
- ARROW, K. J., AND L. HURWICZ (1972): "An Optimality Criterion for Decision-Making under Ignorance," in Uncertainty and Expectations in Economics: Essays in Honour of G. L. S. Shackle, ed. by C. F. Carter, and J. L. Ford, pp. 1–11. Blackwell, Oxford.

- BIRKHOFF, G., AND J. VON NEUMANN (1936): "The Logic of Quantum Mechanics," Annals of Mathematics, 37, 823–843.
- BLUME, L., A. BRANDENBURGER, AND E. DEKEL (1991): "Lexicographic Probabilities and Choice Under Uncertainty," *Econometrica*, 59, 61–79.
- BOLKER, E. D. (1966): "Functions Resembling Quotients of Measures," Transactions of the American Mathematial Society, 124, 292–312.
- (1967): "A Simultaneous Axiomatization of Utility and Subjective Probability," *Philosophy of Science*, 34, 333–340.
- BROOME, J. (1990): "Bolker–Jeffrey Expected Utility Theory and Axiomatic Utilitarianism," Review of Economic Studies, 57, 477–502.
- DEKEL, E. (1986): "An Axiomatic Characterization of Preferences under Uncertainty: Weakening the Independence Axiom," *Journal of Economic Theory*, 40, 304–318.
- DEKEL, E., B. L. LIPMAN, AND A. RUSTICHINI (2001): "Representing Preferences with a Unique Subjective State Space," *Econometrica*, 69, 891–934.
- DEKEL, E., B. L. LIPMAN, A. RUSTICHINI, AND T. SARVER (2005): "Representing Preference with a Unique Subjective State Space: Corrigendum," Working paper, Boston University.
- DUNFORD, N., AND J. T. SCHWARTZ (1957): Linear Operators: Part I. Interscience, New York.
- ELLSBERG, D. (1961): "Risk, Ambiguity, and the Savage Axioms," Quarterly Journal of Economics, 75, 643-669.
- EPSTEIN, L. G. (1999): "A Definition of Uncertainty Aversion," Review of Economic Studies, 66, 579-608.
- EPSTEIN, L. G., AND M. MARINACCI (2005): "Coarse Contingencies," Working paper, University of Rochester and Università di Torino.
- EPSTEIN, L. G., AND J. ZHANG (2001): "Subjective Probabilities on Subjectively Unambiguous Events," *Econometrica*, 69(2), 265–306.
- ERGIN, H., AND F. GUL (2004): "A Subjective Theory of Compound Lotteries," Working paper, Princeton University.
- GAJDOS, T., J.-M. TALLON, AND J.-C. VERGNAUD (2004a): "Coping with Imprecise Information: A Decision Theoretic Approach," Working paper, CNRS–CREST and CNRS–EUREQua.
- (2004b): "Decision Making with Imprecise Probabilistic Information," *Journal of Mathematical Economics*, 40, 647–681.
- GHIRARDATO, P. (2001): "Coping with Ignorance: Unforeseen Contingencies and Non-Additive Uncertainty," Economic Theory, 17, 247–276.
- GHIRARDATO, P., F. MACCHERONI, AND M. MARINACCI (2004): "Differentiatin Ambiguity and Ambiguity Attitude," *Journal of Economic Theory*, 118, 133–173.
- GHIRARDATO, P., AND M. MARINACCI (2002): "Ambiguity Made Precise: A Comparative Foundation," Journal of Economic Theory, 102, 251–289.
- GILBOA, I., AND D. SCHMEIDLER (1989): "Maxmin Expected Utility with a Non–Unique Prior," Journal of Mathematical Economics, 18, 141–153.
- GUL, F., AND W. PESENDORFER (2001): "Temptation and Self-Control," Econometrica, 69, 1403-1435.
- HALEVY, Y. (2004): "Ellsberg Revisited: An Experimental Study," Working paper, University of British Columbia.
- HALEVY, Y., AND V. FELTKAMP (2005): "A Bayesian Approach to Uncertainty Aversion," *Review of Economic Studies*, 72, 449–466.
- HALMOS, P. R. (1974): Measure Theory. Springer, New York, Reprint.
- HAYASHI, T. (2003): "Information, Subjective Belief and Preference," Working paper, University of Rochester.
- JAFFRAY, J.-Y. (1989): "Linear Utility Theory for Belief Functions," Operations Research Letters, 8, 107–112.
- JAFFRAY, J.-Y., AND M. JELEVA (2004): "How to Deal with Partially Analyzed Acts? A Proposal," Working paper, University Paris 6 and University of Nantes.

JEFFREY, R. C. (1965): The Logic of Decision. McGraw-Hill, New York.

KLIBANOFF, P., M. MARINACCI, AND S. MUKERJI (2005): "A Smooth Model of Decision Making under Ambiguity," *Econometrica*, 73, 1849–1892.

- KRANTZ, D. H., R. D. LUCE, P. SUPPES, AND A. TVERSKY (1971): Foundations of Measurement, vol. 1. Academic Press, New York.
- KREPS, D. M. (1979): "A Representation Theorem for 'Preference for Flexibility'," Econometrica, 47, 565–576.
- MACHINA, M. J., AND D. SCHMEIDLER (1992): "A More Robust Definition of Subjective Probability," *Econometrica*, 60, 745–780.
- NAU, R. F. (2003): "Uncertainty Aversion with Second–Order Utilities and Probabilities," Working paper, Fuqua School of Business, Duke University.

NEHRING, K. (1999): "Preference for Flexibility in a Savage Framework," Econometrica, 67, 101–119.

(2001): "Ambiguity in the Context of Probabilistic Beliefs," Working paper, University of California, Davis.

OLSZEWSKI, W. B. (2003): "Preferences over Sets of Lotteries," Working paper, Northwestern University.

SAVAGE, L. J. (1954): The Foundations of Statistics. Wiley, New York.

SEGAL, U. (1987): "The Ellsberg Paradox and Risk Aversion: An Anticipated Utility Approach," International Economic Review, 58, 349–377.

(1990): "Two-Stage Lotteries without the Reduction Axion," *Econometrica*, 58, 349–377.

SINISCALCHI, M. (forthcoming): "A Behavioral Characterization of Plausible Priors," Journal of Economic Theory.

- STINCHCOMBE, M. B. (2003): "Choice and Games with Ambiguity as Sets of Probabilities," Working paper, University of Texas, Austin.
- SUPPES, P. (1966): "The Probabilistic Argument for a Non-Classical Logic of Quantum Mechanics," Philosophy of Science, 33, 14–21.
- VON NEUMANN, J. (1955): Mathematical Foundations of Quantum Mechanics. Princeton University Press, Princeton, translated by R. T. Beyer.
- VON NEUMANN, J., AND O. MORGENSTERN (1944): Theory of Games and Economic Behavior. Princeton University Press, Princeton.
- WANG, T. (2003): "A Class of Multi-Prior Preferences," Working paper, University of British Columbia.
- ZHANG, J. (1999): "Qualitative Probabilities on λ -Systems," Mathematical Social Sciences, 38, 11–20.